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Dynamic optimal execution in a mixed-market-impact Hawkes price model

Aurélien Alfonsi, Pierre Blanc*

June 9, 2015

Abstract

We study a linear price impact model including other liquidity takers, whose flow of orders follows a Hawkes process. The optimal execution problem is solved explicitly in this context, and the closed-formula optimal strategy describes in particular how one should react to the orders of other traders. This result enables us to discuss the viability of the market. It is shown that Poissonian arrivals of orders lead to quite robust Price Manipulation Strategies in the sense of Huberman and Stanzl [29]. Instead, a particular set of conditions on the Hawkes model balances the self-excitation of the order flow with the resilience of the price, excludes Price Manipulation Strategies and gives some market stability.

Keywords: Market Impact Model, Optimal Execution, Hawkes Processes, Market Microstructure, High-frequency Trading, Price Manipulations.

AMS (2010): 91G99, 91B24, 91B26, 60G55, 49J15.

JEL: C02, C61, C62.

1 Introduction

When modeling the price of an asset, we typically distinguish at least three different time scales. At the low-frequency level, the price can often be well approximated by a diffusive process. At the other end, when dealing with very high frequencies, some key features of the Limit Order Book (LOB) dynamics have to be modeled. In between, price impact models consider an intra-day mesoscopic time scale, somewhere between seconds and hours. They usually ignore most of the LOB events (limit orders, cancellations, market orders, etc.) and focus on describing the price impact of the transactions. Their goal is to be more tractable than high-frequency models and to bring quantitative results on practical issues such as optimal execution strategies. The usual setup is well-described in Gatheral [22], who defines the price process S as

$$S_t = S_0 + \int_0^t f(\dot{x}_s)G(t-s)ds + \int_0^t \sigma dZ_s,$$

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where \dot{x}_s is the rate of trading of the liquidating agent at time $s < t$, $f(v)$ represents the instantaneous price impact of an agent trading at speed v , G is called a “decay kernel” and Z is a noise process. The quantity $f(v)G(+\infty)$ is usually called the “permanent impact”, $f(v)G(0)$ the “immediate impact” and $f(v)[G(0^+) - G(+\infty)]$ the “transient impact”. The pioneering price impact models of Bertsimas and Lo [10] and Almgren and Chriss [5] consider a linear impact, with an immediate and a permanent part (which corresponds to $f(v) = \alpha v$, $G(0) > 0$, $G(0^+) = G(+\infty) > 0$ with the previous notations). These models ignore the transient part of the impact which is due to the resilience of the market and cannot be neglected when trading frequently. For that purpose, Obizhaeva and Wang [32] have considered a model that includes in addition a linear transient impact that decays exponentially (i.e. $f(v) = \alpha v$, $G(u) = \lambda + (1 - \lambda) \exp(-\rho u)$, $0 \leq \lambda \leq 1$, $\rho > 0$). However, empirical evidence on market data shows that the price impact is not linear but rather concave, see e.g. Potters and Bouchaud [33], Eisler et al. [17], Mastromatteo, Tóth and Bouchaud [31], Donier [16] and more recently, Farmer, Gerig, Lillo and Waelbroeck [19]. Extensions or alternatives to the Obizhaeva and Wang model that include non-linear price impact have been proposed by Alfonsi, Fruth and Schied [2], Predoiu, Shaikhet and Shreve [34], Gatheral [22] and Guéant [24] to mention a few. Similarly, the exponential decay of the transient impact is not truly observed on market data, and one should consider more general decay kernels. Alfonsi, Schied and Slynko [4] and Gatheral, Schied and Slynko [23] consider the extension of the Obizhaeva and Wang model when the transient impact has a general decay kernel. Another simplification made by these models is that they generally assume that when the liquidating trader is passive, the price moves according to a continuous martingale, that sums up the impact of all the orders issued by other participants. However, if one wants to use these models at a higher frequency, they would naturally wonder how these orders (at least the largest ones) can be taken into account in the strategy, and if the martingale hypothesis for the price can be relaxed. This is one of the contributions of the present paper.

On the other hand, high-frequency price models aim at reproducing some statistical observations made on market data such as the autocorrelation in the signs of trades, the volatility clustering effect, the high-frequency resilience of the price, etc., and to obtain low-frequency asymptotics that are consistent with continuous diffusions. At very high frequencies, one then has to describe LOB dynamics, or a part of it. Such models have been proposed by Abergel and Jedidi [1], Huang, Lehalle and Rosenbaum [28], Cont and de Larrard [13], Garèche et al. [21], among others. However, as stressed in [13], LOB events are much more frequent than price moves. Thus, it may be relevant to model the price at the slightly lower frequency of midpoint price changes. For example, Robert and Rosenbaum [35] have proposed a model based on a diffusion with uncertainty zones that trigger the price changes. Recently, Bacry et al. [6] presented a tick-by-tick price model based on Hawkes processes, that reproduces well some empirical facts of market data. This model has then been enriched by Bacry and Muzy [8] to describe jointly the order flow and the price moves. In fact, there is a very recent and active literature that focuses on the use of mutually exciting Hawkes processes in high-frequency price models. Without being exhaustive, we mention here the works of Da Fonseca and Zaatour [14], Zheng, Roueff and Abergel [38], Filimonov and Sornette [20] and Hardiman, Bercot and Bouchaud [25]. Asymptotic and low-frequency behaviour of such models has been investigated recently by Bacry et al. [7] and Jaisson and Rosenbaum [30].

The present paper is a contribution to this also mutually exciting literature. Its main goal is to make a bridge between high-frequency price models and optimal execution frameworks. On the one hand, Hawkes processes seem to be rich enough to describe satisfactorily the flow of market orders. On the other hand, price impact models are tractable and well-designed to calculate trading costs. The aim of our model is to grasp these two features. Thus, we consider an Obizhaeva and Wang framework where market buy and sell orders issued by other traders are modeled through Hawkes processes. This enables us to make quantitative calculations and to solve the optimal execution problem explicitly. We obtain a necessary and sufficient condition on the parameters of the Hawkes model to rule out Price Manipulation Strategies that can be seen as high-frequency arbitrages. Interestingly, we also show that modeling the order flow with a Poisson process necessarily leads to those arbitrages.

The paper is organized as follows. In Section 2, we set up the model and present a general criterion to exclude Price Manipulation Strategies. Section 3 summarizes our main results. Section 4 gives the solution of the

optimal execution problem along with several comments and insights on the optimal strategy. Eventually, we analyze the existence of Price Manipulation Strategies in our model in Section 5 and give the conditions under which they are impossible. Cumbersome explicit formulas and technical proofs are gathered in the Appendix.

2 Model setup and the optimal execution problem

2.1 General price model

We start by describing the price model itself, without considering the execution problem. We consider a single asset and denote by P_t its price at time t . We assume that we can write it as the sum of a “fundamental price” component S_t and a “mesoscopic price deviation” D_t :

$$P_t = \underbrace{S_t}_{\text{fundamental price}} + \underbrace{D_t}_{\text{mesoscopic price deviation}}. \quad (1)$$

Typically, these quantities are respectively related to the permanent and the transient impact of the market orders. We now specify the model and consider the framework of Obizhaeva and Wang [32] where these impacts are linear. Let N_t be the sum of the signed volumes of past market orders on the book between time 0 and time t . By convention, a buy order is counted positively in N while a sell order makes N decrease, and we assume besides that N is a càdlàg (right continuous with left limits) process. We assume that an order modifies the price proportionally to its size, which would correspond to a block-shaped limit order book. A proportion $\nu \in [0, 1]$ of the price impact is permanent, while the remaining proportion $1 - \nu$ is transient with an exponential decay of speed $\rho > 0$. This mean-reversion effect can be seen as the feedback of market makers, who affect the price using limit orders and cancellations. Namely, we consider the following dynamics for S and D :

$$\begin{aligned} dS_t &= \frac{\nu}{q} \underbrace{dN_t}_{\text{market orders}} \\ dD_t &= \underbrace{-\rho D_t dt}_{\text{market resilience}} + \frac{1-\nu}{q} \underbrace{dN_t}_{\text{market orders}}, \end{aligned}$$

with $q > 0$. One should note that in this model, the variations in the fundamental value of the asset are revealed in its price through the process S . Indeed, we assume that the impact of each incoming market order, modeled through the process N , contains of proportion ν of “real” or “exogenous” information, and that the remaining proportion $1 - \nu$ is of endogenous origin and will vanish over time.

Remark 2.1. *This model assumes a linear price impact with an exponential resilience. As mentioned in the introduction, these assumptions are challenged by empirical facts, and it would be for sure interesting and relevant to enrich the model by considering a non linear price impact and a more general decay of the impact. However, the new feature of the model with respect to the literature on optimal execution is to add a flow of market orders issued by other traders. This is why we afford to make these simplifying assumptions that give analytical tractability, which is important to calculate the optimal execution strategy in real time. Thus, the model is meant to constitute a first step in dynamic optimal execution with the price driven by point processes, and we plan to confront it to market data in a future work.*

As usual, we consider $(\Omega, \mathcal{F}, \mathbb{P})$ a probability space where \mathbb{P} weights the probability of the market events. We assume that the process $(N_t)_{t \geq 0}$ has bounded variation and is square integrable, i.e. $\sup_{s \in [0, t]} \mathbb{E}[N_s^2] < \infty$ for any $t \geq 0$, and we define $(\mathcal{F}_t)_{t \geq 0}$ the natural filtration of N , $\mathcal{F}_t = \sigma(N_s, s \leq t)$ for $t \geq 0$. We will specify in Section 2.3 which dynamics we consider for N in this paper.

2.2 Optimal execution framework

We now consider a particular trader who wants to buy or sell a given quantity of assets on the time interval $[0, T]$. Through the paper, we will call this trader the “strategic trader” to make the distinction between his market orders and all the other market orders, that are described by N . We will denote by X_t the number of assets owned by the strategic trader at time t . We assume that the process is (\mathcal{F}_t) -adapted, with bounded variation and càglàd (left continuous with right limits) which means that the strategic trader observes all the information available on the market, and that he can react instantly to the market orders issued by other traders. Besides, a strategy that liquidates x_0 assets on $[0, T]$ should satisfy $X_0 = x_0$ and $X_{T+} = 0$: $x_0 > 0$ (resp. $x_0 < 0$) corresponds to a buy (resp. sell) program.

Definition 2.1. *A liquidating strategy X for the position $x_0 \in \mathbb{R}$ on $[0, T]$ is admissible if it is (\mathcal{F}_t) -adapted, càglàd, square integrable, with bounded variation and such that $X_0 = x_0$ and $X_{T+} = 0$, a.s.*

Remark 2.2. *An admissible strategy X has a countable set \mathcal{D}_X of times of discontinuity on $[0, T]$, and can have a non-zero continuous part $X_t^c = X_t - \sum_{\tau \in \mathcal{D}_X \cap [0, t)} (X_{\tau+} - X_\tau)$, $t \in [0, T]$.*

One then has to specify how the strategic trader modifies the price, as well as the cost induced by his trading strategy. Again, we will consider the Obizhaeva and Wang model [32] with the same price impact as above. However, we let the possibility that the proportion $\epsilon \in [0, 1]$ of permanent impact of the strategic trader could be different from the one of the other traders, which we note $\nu \in [0, 1]$. Of course, a reasonable choice would be to set $\epsilon = \nu$ to consider all orders equally, but the model allows for more generality. We then assume the following dynamics

$$dS_t = \frac{1}{q} (\nu dN_t + \epsilon dX_t), \quad (2)$$

$$dD_t = -\rho D_t dt + \frac{1}{q} ((1 - \nu)dN_t + (1 - \epsilon)dX_t). \quad (3)$$

With the assumptions on N and X , the price processes P , S and D have left and right limits. More precisely, in case of discontinuity at time t , (2) and (3) have to be read here as follows

$$\begin{aligned} S_t - S_{t-} &= \frac{\nu}{q} (N_t - N_{t-}), \quad S_{t+} - S_t = \frac{\epsilon}{q} (X_{t+} - X_t), \\ D_t - D_{t-} &= \frac{1 - \nu}{q} (N_t - N_{t-}), \quad D_{t+} - D_t = \frac{1 - \epsilon}{q} (X_{t+} - X_t). \end{aligned}$$

For the sake of tractability only, we make the assumption of a block-shaped Limit Order Book. Thus (see [32]), when the strategic trader places at time t an order of size $v \in \mathbb{R}$ ($v > 0$ for a buy order and $v < 0$ for a sell order), it has the following cost

$$\pi_t(v) = \int_0^v \left[P_t + \frac{1}{q} y \right] dy = \underbrace{P_t v}_{\text{cost at the current price}} + \underbrace{\frac{v^2}{2q}}_{\text{impact cost}}.$$

Since $P_{t+} = P_t + \frac{v}{q}$, this cost amounts to trade all the assets at the average price $(P_t + P_{t+})/2$. We stress here that if an order has just occurred, i.e. $N_t - N_{t-} \neq 0$, the value of P_t is different from P_{t-} and takes into account the price impact of this order. Therefore, the cost of an admissible strategy X is given by

$$\begin{aligned} C(X) &= \int_{[0, T)} P_u dX_u + \frac{1}{2q} \sum_{\tau \in \mathcal{D}_X \cap [0, T)} (\Delta X_\tau)^2 - P_T X_T + \frac{1}{2q} X_T^2 \\ &= \int_{[0, T)} P_u dX_u^c + \sum_{\tau \in \mathcal{D}_X \cap [0, T)} P_\tau (\Delta X_\tau) + \frac{1}{2q} \sum_{\tau \in \mathcal{D}_X \cap [0, T)} (\Delta X_\tau)^2 - P_T X_T + \frac{1}{2q} X_T^2, \end{aligned} \quad (4)$$

since at time T all the remaining assets have to be liquidated. Here, the sum brings on the countable times of discontinuity \mathcal{D}_X of X , and the jumps $\Delta X_\tau = X_{\tau+} - X_\tau \neq 0$ for $\tau \in \mathcal{D}_X$. We note that all the terms involved in the cost function are integrable, thanks to the assumption on the square integrability of X and N .

Remark 2.3. *With the initial market price P_0 taken as a reference, $-P_0 \times x_0$ is the mark-to-market liquidation cost. Thus, $C(X) + P_0 \times x_0$ can be seen as an additional liquidity cost of it is positive. If it is negative, its absolute value can be seen as the gain associated to the strategy X .*

Remark 2.4. *The cost defined by (4) in the price model (1), (2) and (3) is a deterministic function of $(X_t)_{t \in [0, T]}$, $(N_t)_{t \in [0, T]}$, S_0 , D_0 and the parameters q , ν , and ϵ . In this remark, we denote by $C(X, N, S_0, D_0, q)$ this function when ν and ϵ are given. From (2), (3) and (4), we have the straightforward property*

$$C(X, N, S_0, D_0, q) = C(-X, -N, -S_0, -D_0, q). \quad (5)$$

Observing that $qC(X) = \int_{[0, T]} qP_u dX_u + \frac{1}{2} \sum_{0 \leq \tau < T} (\Delta X_\tau)^2 - (qP_T)X_T + \frac{1}{2}(X_T)^2$, and remarking that qS and qD satisfy (2) and (3) with $q = 1$, we also get

$$qC(X, N, S_0, D_0, q) = C(X, N, qS_0, qD_0, 1). \quad (6)$$

Remark 2.5. *Since X is a càglàd process and N is a càdlàg process, we will have to work with làdlàg (with finite right-hand and left-hand limits) processes. When Z is a làdlàg process, we set $\Delta^- Z_t = Z_t - Z_{t-}$ and $\Delta^+ Z_t = Z_{t+} - Z_t$ the left and right jumps of Z , and $Z_t^c = Z_t - \sum_{0 \leq \tau < t} \Delta^+ Z_\tau - \sum_{0 < \tau \leq t} \Delta^- Z_\tau$ the continuous part of Z . We also set $\Delta Z_t = Z_{t+} - Z_{t-}$ and use the shorthand notation $dZ_t = dZ_t^c + \Delta Z_t$. If $dZ_t = d\tilde{Z}_t$ for some other làdlàg process \tilde{Z} , this means that $dZ_t^c = d\tilde{Z}_t^c$ and $\Delta Z_t = \Delta \tilde{Z}_t$. In particular, when Z is càdlàg and \tilde{Z} is càglàd, this means that $Z_t - Z_{t-} = \tilde{Z}_{t+} - \tilde{Z}_t$ at the jump times.*

Then, the optimal execution problem consists in finding an admissible strategy X that minimizes the expected cost $\mathbb{E}[C(X)]$ for a given initial position $x_0 \in \mathbb{R}$. This problem for $x_0 = 0$ is directly related to the existence of Price Manipulation Strategies as defined below.

Definition 2.2. *A Price Manipulation Strategy (PMS) in the sense of Huberman and Stanzl [29] is an admissible strategy X such that $X_0 = X_{T+} = 0$ a.s. for some $T > 0$ and $\mathbb{E}[C(X)] < 0$.*

We have the following result that gives a necessary and sufficient condition to exclude PMS.

Theorem 2.1. *The model does not admit PMS if, and only if the process P is a (\mathcal{F}_t) -martingale when $X \equiv 0$. In this case, the optimal strategy X^{OW} is the same as in the Obizhaeva and Wang [32] model. It is given by*

$$\Delta X_0^{\text{OW}} = -\frac{x_0}{2 + \rho T}, \quad \Delta X_T^{\text{OW}} = -\frac{x_0}{2 + \rho T}, \quad dX_t^{\text{OW}} = -\rho \frac{x_0}{2 + \rho T} dt \quad \text{for } t \in (0, T), \quad (7)$$

and has the expected cost $\mathbb{E}[C(X^{\text{OW}})] = -P_0 x_0 + \left[\frac{1-\epsilon}{2+\rho(T-t)} + \frac{\epsilon}{2} \right] x_0^2 / q$.

This theorem is proved in Appendix C. Similar results are standard in financial mathematics, but to the best of our knowledge, it has not yet been formulated as such in the literature in a context with price impact and with respect to the notion of Price Manipulation Strategies. In usual optimal execution frameworks, the unaffected price is assumed *a priori* to be a martingale, which is not the case here. Note that if P is a martingale, the optimal strategy is very robust in the sense that it does not depend on N , and is therefore the same as the one in the Obizhaeva and Wang model [32] that corresponds to $N \equiv 0$ and $D_0 = 0$. In fact, it does not depend either on ϵ and ν , and only depends on ρ .

Theorem 2.1 indicates that suitable models for the order flow N should be such that P is, roughly speaking, close to a martingale when the strategic trader is absent, so that arbitrage opportunities are short-lived and

not too visible. This raises at least three questions. Which “simple” processes N can lead to a martingale price P ? Can we characterize the optimal strategy when P is not a martingale? In particular, in the latter case, how does the optimal strategy take the market orders issued by other participants into account? In this paper, we study these questions when N follows a Hawkes process.

Remark 2.6. *The model can be generalized by adding a càdlàg (\mathcal{F}_t) -martingale S^0 to the price process P , i.e. if we replace (1) by $P_t = S_t + D_t + S_t^0$, with $S_0^0 = 0$. This does not change the optimal execution problem since, using an integration by parts, S^0 adds the following term to the cost*

$$\begin{aligned} \int_{[0,T)} S_t^0 dX_t - S_T^0 X_T &= S_T^0 X_T - S_0^0 X_0 - \int_{[0,T)} X_t dS_t^0 - S_T^0 X_T \\ &= - \int_{[0,T)} X_t dS_t^0, \end{aligned}$$

which has a zero expected value from the martingale property. Let us note that there is no covariation between the processes X and S_0 since they do not jump simultaneously and X has bounded variations.

Remark 2.7. *Similarly, when N is a càdlàg (\mathcal{F}_t) -martingale and X is an admissible liquidating strategy for $X_0 = x_0$, we have*

$$\mathbb{E}[C(X)] = \mathbb{E} \left[\int_{[0,T)} D_u dX_u + \frac{1-\epsilon}{2q} \sum_{0 \leq \tau < T} (\Delta X_\tau)^2 - D_T X_T + \frac{1-\epsilon}{2q} X_T^2 \right] + \frac{\epsilon}{2q} x_0^2,$$

since $x_0^2 = \int_{[0,T+]} d[(X_t - X_0)^2] = 2 \int_{[0,T)} (X_u - X_0) dX_u + \sum_{0 \leq \tau < T} (\Delta X_\tau)^2 - 2(X_T - X_0)X_T + X_T^2$. When $\epsilon \in [0, 1)$, we set $X_t^\epsilon = (1 - \epsilon)X_t$ and get

$$\mathbb{E}[C(X)] = \frac{1}{q(1-\epsilon)} \mathbb{E} \left[\int_{[0,T)} q D_u d(X_u^\epsilon) + \frac{1}{2} \sum_{0 \leq \tau < T} (\Delta X_\tau^\epsilon)^2 - q D_T X_T^\epsilon + \frac{1}{2} (X_T^\epsilon)^2 \right] + \frac{\epsilon}{2q} x_0^2. \quad (8)$$

Therefore, X is optimal if, and only if X^ϵ is optimal in the model with $\epsilon = \nu = 0$, $q = 1$ and an incoming flow of market orders equal to $(1 - \nu)N$.

2.3 The MIH model

2.3.1 Definitions and notations

We introduce the MIH (Mixed-market-Impact Hawkes) price model, where

$$N_t = N_t^+ - N_t^-,$$

the process (N^+, N^-) being a symmetric two-dimensional marked Hawkes process of intensity (κ^+, κ^-) . The process $(N^+, N^-, \kappa^+, \kappa^-)$ is càdlàg and jumps when N jumps. We note $n^+(dt, dv)$ and $n^-(dt, dv)$ the Poisson measures on $\mathbb{R}^+ \times \mathbb{R}^+$ associated to N^+ and N^- respectively, where the variable v stands for the amplitudes of the jumps, i.e. the volumes of incoming market orders. We restrain to the case of i.i.d. unpredictable marks of common law μ on \mathbb{R}^+ , i.e. for any $A \in \mathcal{B}(\mathbb{R}^+)$ and $t \geq 0$,

$$\kappa_t^\pm \mu(A) = \lim_{h \rightarrow 0^+} \frac{1}{h} \mathbb{E}[n^\pm([t, t+h], A) | \mathcal{F}_t],$$

where $\mathcal{F}_t = \sigma(N_u^+, N_u^-, u \leq t) = \sigma(N_u, u \leq t)$ as defined earlier. In other words, at time t , the conditional instantaneous jump intensity of N^\pm is given by κ_t^\pm , and the amplitudes of the jumps are i.i.d. variables of law μ which are independent from the past, i.e. from \mathcal{F}_{t-} . We also define

$$m_k = \int_{\mathbb{R}^+} v^k \mu(dv), \quad k \in \mathbb{N},$$

assuming moreover that $m_2 < \infty$. We choose the Hawkes kernel to be the exponential $t \mapsto \exp(-\beta t)$, $\beta \geq 0$, so that $(N^+, N^-, \kappa^+, \kappa^-)$ is Markovian. Thus, we set

$$\begin{pmatrix} \kappa_t^+ \\ \kappa_t^- \end{pmatrix} = \begin{pmatrix} \kappa_\infty \\ \kappa_\infty \end{pmatrix} + \left[\begin{pmatrix} \kappa_0^+ \\ \kappa_0^- \end{pmatrix} - \begin{pmatrix} \kappa_\infty \\ \kappa_\infty \end{pmatrix} \right] \exp(-\beta t) + \int_0^t \int_{(\mathbb{R}^+)^2} \exp(-\beta(t-u)) \begin{pmatrix} \varphi_s(v^+/m_1) & \varphi_c(v^-/m_1) \\ \varphi_c(v^+/m_1) & \varphi_s(v^-/m_1) \end{pmatrix} \cdot \begin{pmatrix} n^+(du, dv^+) \\ n^-(du, dv^-) \end{pmatrix},$$

where $\kappa_\infty \geq 0$ is the common baseline intensity of N^+ and N^- , and $\varphi_s, \varphi_c : \mathbb{R}^+ \rightarrow \mathbb{R}^+$ are measurable positive functions that satisfy

$$\iota_s := \int_{\mathbb{R}^+} \varphi_s(v/m_1) \mu(dv) < \infty \quad , \quad \iota_c := \int_{\mathbb{R}^+} \varphi_c(v/m_1) \mu(dv) < \infty.$$

We assume besides that

$$\int_{\mathbb{R}^+} \varphi_s^2(v/m_1) \mu(dv) < \infty, \quad \int_{\mathbb{R}^+} \varphi_c^2(v/m_1) \mu(dv) < \infty$$

to have $\sup_{s \in [0, t]} \mathbb{E}[N_s^2] < \infty$, and we note that this property is automatically satisfied when φ_s and φ_c have a sublinear growth since we have assumed $m_2 < \infty$. From the modeling point of view, we may expect that the functions φ_s and φ_c are nondecreasing: the larger an order is, the more other orders it should trigger. However, we do not need this monotonicity assumption in the mathematical analysis.

Equivalently, in this Markovian setting, the intensities κ_t^+ and κ_t^- follow the dynamics

$$\begin{aligned} d\kappa_t^+ &= -\beta (\kappa_t^+ - \kappa_\infty) dt + \varphi_s(dN_t^+/m_1) + \varphi_c(dN_t^-/m_1), \\ d\kappa_t^- &= -\beta (\kappa_t^- - \kappa_\infty) dt + \varphi_c(dN_t^+/m_1) + \varphi_s(dN_t^-/m_1), \end{aligned} \tag{9}$$

where formally, $\int_0^t \varphi_s(dN_u^+/m_1) = \int_0^t \int_{\mathbb{R}^+} \varphi_s(v/m_1) n^+(du, dv)$ for $t \geq 0$. As pointed out in Hardiman, Bercot and Bouchaud [25] and Bacry and Muzy [8] for instance (in a slightly different context since in our framework, N models market orders only), a power-law Hawkes kernel is more in accordance with market data than an exponential one. It is possible in principle to approximate a completely monotone decaying kernel with a multi-exponential one while preserving a Markovian framework, at the cost of increasing the dimension of the state space, see for example Alfonsi and Schied [3]. This investigation is left for future research.

Note that N^+ and N^- boil down to independent composed Poisson processes in the case $\beta = 0$, $\varphi_s = \varphi_c \equiv 0$. The meaning of the parameters is rather clear: κ^+ and κ^- are mean reverting processes, and ι_s and ι_c respectively describe how a market buy order increases the instantaneous probability of buy (resp. sell) orders. More precisely, ι_s encodes both the splitting of meta-orders, and the fact that participants tend to follow market trends (which is called the herding effect). On the other hand, ι_c describes opportunistic traders that sell (resp. buy) after a sudden rise (resp. fall) of the price. The functions φ_s and φ_c allow respectively the self and cross-excitations in the order flow to depend on the volumes of the orders. For instance, for constant functions $\varphi_s \equiv \iota_s$ and $\varphi_c \equiv \iota_c$, the model boils down to the standard Hawkes model where κ^\pm makes jumps of constant size when N^\pm jumps.

Hawkes processes have been recently used in the literature to model the price. In particular, Bacry et al. [6] consider a similar model where N models all price moves, with $\nu = 1$, $\iota_s = 0$ and deterministic jumps (i.e. μ is a Dirac mass). More recently, Bacry and Muzy [8] have proposed an four-dimensional Hawkes process to model the market buy and sell orders together with the up and down events on the price. In contrast, the model that we study here determines the price impact of an order in function of its size. For the reader who is not accustomed to Hawkes processes, we point the original paper [26], the paper by Embrechts et al. [18] for an overview of multivariate marked Hawkes processes and the book of Daley and Vere-Jones [15] for a more detailed account.

Remark 2.8. *As one can see in Equation (9), the orders of the strategic trader do not impact the jump rates κ^+ and κ^- (there is no dX_t term), as opposed to the market orders issued by other traders. The first reason*

for this modeling choice is tractability. However, it is found empirically by Tóth et al. [37] that the main contribution to the self-excitation of the order flow comes from the splitting effect. Each individual trader tends to post several orders of the same nature (buy or sell) in a row, which creates auto-correlation in the signs of trades, and this effect is significantly stronger than the mutual excitation between different traders. Thus, it is an acceptable approximation to neglect the excitation coming from the orders of the strategic trader. Of course, it would be nice to find in the future a tractable model that gives a unified framework for the mutual excitation that considers equally all the market orders.

2.3.2 Stationarity and low-frequency asymptotics of the MIH model

Up to now, we have presented the MIH model without assuming stationarity. In most models featuring Hawkes processes, stationarity is an *a priori* assumption, but here, we do not need it to derive the optimal strategy. However, if one wishes to use the MIH model with constant parameters on a large time period, it may be reasonable to consider parameters that satisfy stationarity. This is why we present here a few results that are standard in the literature of Hawkes processes.

We consider the MIH model when the strategic trader is absent, i.e. $X \equiv 0$.

Proposition 2.1. *The process (κ_t^+, κ_t^-) converges to a stationary law if, and only if $\iota_s + \iota_c < \beta$.*

Proof. We can apply the results of the existing literature on marked Hawkes processes with unpredictable marks (for instance Hawkes and Oakes [27], Brémaud and Massoulié [12] or Daley and Vere-Jones [15]) to obtain that (κ_t^+, κ_t^-) converges to a stationary law if the largest eigenvalue of

$$\int_{\mathbb{R}^+ \times \mathbb{R}^+} \exp(-\beta t) \begin{pmatrix} \varphi_s(v/m_1) & \varphi_c(v/m_1) \\ \varphi_c(v/m_1) & \varphi_s(v/m_1) \end{pmatrix} dt \mu(dv) = \frac{1}{\beta} \begin{pmatrix} \iota_s & \iota_c \\ \iota_c & \iota_s \end{pmatrix}$$

is strictly below unity. Conversely, if $\iota_s + \iota_c \geq \beta$, we have

$$\frac{d}{dt} \mathbb{E}[\kappa_t^+ + \kappa_t^-] = 2\beta\kappa_\infty + (\iota_s + \iota_c - \beta)\mathbb{E}[\kappa_t^+ + \kappa_t^-] \geq 2\beta\kappa_\infty$$

and the process cannot be stationary. □

We now study the low-frequency asymptotics of the price process P in the MIH model. We consider the sequence $P_t^{(n)} = P_{nt}/\sqrt{n}$ for $n \geq 1$. We have $P_t^{(n)} = S_t^{(n)} + D_t^{(n)}$, where we also set $S_t^{(n)} = S_{nt}/\sqrt{n}$ and $D_t^{(n)} = D_{nt}/\sqrt{n}$. To study the behaviour of $D^{(n)}$, we need the following lemma.

Lemma 2.1. *When $\iota_s + \iota_c < \beta$, the expectation $\mathbb{E}[D_t^2]$ converges to a finite positive value as $t \rightarrow +\infty$.*

The proof of this lemma is rather straightforward. We just have to calculate $\mathbb{E}[\delta_t^2]$, $\mathbb{E}[\delta_t D_t]$ and $\mathbb{E}[D_t^2]$ and check that these expectations converge when $\iota_s + \iota_c < \beta$. This result implies that $(D_{t_1}^{(n)}, \dots, D_{t_k}^{(n)})$ converges to zero for the L^2 norm for any $0 \leq t_1 \leq \dots \leq t_k$. This gives that the process $D^{(n)}$ converges to zero.

We thus focus on the convergence of $S_t^{(n)} = \frac{\nu}{q} \frac{N_{nt}^+ - N_{nt}^-}{\sqrt{n}}$. If the jumps of N are bounded, i.e. μ has bounded support, and $v \mapsto \varphi_s(v/m_1)$ and $v \mapsto \varphi_c(v/m_1)$ are bounded on the support of μ (which are reasonable assumptions in practice), a straightforward adaptation of Corollary 1 of Bacry et al. [7] gives the convergence in law of $S^{(n)}$ to a non-standard Brownian motion with zero drift.

3 Main results

Now that the whole framework is set up, we present the main results of the present paper.

- The optimal execution problem can be solved explicitly in the MIH model and the optimal strategy has still a quite simple form, see Theorem 4.1. Of course, this result relies on the assumptions of linear price impact and exponential decay kernel, which are not in accordance with empirical facts, see for example Potters and Bouchaud [33] and Bouchaud et al. [11]. We mention here that it would be possible to keep an affine structure of the optimal strategy by considering complete monotone decay kernels as in Alfonsi and Schied [3]. However, we believe that the optimal strategy is interesting at least from a qualitative point of view, since it gives clear insights on how to react optimally to observed market orders and on the role of the different parameters of the model.
- Price Manipulation Strategies necessarily appear when the flow of market orders is Poissonian, and they are rather robust in the sense that they can be implemented without knowing the model parameters. Namely, the strategy which consists in trading instantly a small proportion of the volume of each incoming market order in the opposite direction is profitable on average, see Proposition 5.2. This justifies to consider more elaborate dynamics for the order arrivals.
- Even in a non-Poissonian MIH setup, Price Manipulation Strategies can arise. Depending on the parameters of the model and on the size of each observed market order, one should either trade instantly in the opposite direction to take market resilience into account, or in the same direction to take advantage of the self-excitation property of Hawkes processes. However, our framework allows for a specific equilibrium to take place, that we call the Mixed-market-Impact Hawkes Martingale (MIHM) model, where PMS disappear.
- In the MIHM model, one has in particular $\iota_s > \iota_c$, $\nu < 1$ and $\beta = \rho$, and the self-excitation property of the order flow exactly compensates the price resilience induced by market makers. The resulting price process is a martingale even at high frequencies, and in this case we find that the optimal strategy and cost function are those of Obizhaeva and Wang [32]. The conditions of this model imply that if $\iota_c = 0$, the norm ι_s/β of the Hawkes kernel that symbolizes the endogeneity ratio of the market, see Filimonov and Sornette [20], should be equal to $1 - \nu$, i.e. the proportion of market impact which is transient.
- The fact of reacting to the market orders of other traders with instantaneous market orders can trigger chain reactions and lead to market instability. We show that in the MIH framework, the conditions under which it is profitable for the strategic trader to react instantaneously to other trades are quite equivalent to the existence of PMS. Although the model is clearly a simplified view of the market, it is remarkable to obtain in this case such a clear connection between market stability and free profits. It would be interesting for further research to investigate if this conclusion still holds in a more general model.

4 The optimal strategy

We need to introduce some notations to present the main results on the optimal execution. Instead of working with κ_t^+ and κ_t^- , we will rather use $\delta_t = \kappa_t^+ - \kappa_t^-$ and $\Sigma_t = \kappa_t^+ + \kappa_t^-$ that satisfy from (9)

$$d\delta_t = -\beta \delta_t dt + dI_t \quad , \quad d\Sigma_t = -\beta (\Sigma_t - 2\kappa_\infty) dt + d\bar{I}_t, \quad (10)$$

where

$$\begin{aligned} I_t &= \int_0^t [(\varphi_s - \varphi_c)(dN_u^+/m_1) - (\varphi_s - \varphi_c)(dN_u^-/m_1)] , \\ \bar{I}_t &= \int_0^t [(\varphi_s + \varphi_c)(dN_u^+/m_1) + (\varphi_s + \varphi_c)(dN_u^-/m_1)] . \end{aligned} \quad (11)$$

The processes I and \bar{I} are càdlàg processes which describe intensity jumps, and their jump times are those of N . In the standard Hawkes framework where φ_s and φ_c are constant, one has $\varphi_s \equiv \iota_s$ and $\varphi_c \equiv \iota_c$, and when N jumps, I jumps of $(\iota_s - \iota_c) \text{sgn}(\Delta N_t)$ and \bar{I} of $\iota_s + \iota_c$.

We note $(\tau_i)_{i \geq 1}$ the ordered random jump times of N and set $\tau_0 = 0$. For $t \in [0, T]$, we also note χ_t the total number of jumps of I that occurred between time 0 and time t . From (10), we have

$$\delta_t = \delta_0 \exp(-\beta t) + \sum_{l=1}^{\chi_t} \exp(-\beta(t - \tau_l)) \Delta I_{\tau_l} = \delta_0 \exp(-\beta t) + \exp(-\beta t) \Theta_{\chi_t},$$

where we define $\Theta_0 = 0$ and

$$\Theta_i = \sum_{l=1}^i \exp(\beta \tau_l) \Delta I_{\tau_l} = \sum_{0 < \tau \leq \tau_i} \exp(\beta \tau) \Delta I_{\tau}, \quad i \geq 1.$$

For $i \geq 0$ and $t \in [\tau_i, \tau_{i+1})$, we obtain that $\delta_t \exp(\beta t) = \delta_0 + \Theta_i$ only depends on t through the integer $i = \chi_t$. We introduce the useful quantities

$$\alpha = \iota_s - \iota_c, \quad \eta = \beta - \alpha,$$

and the two continuously differentiable functions $\zeta, \omega : \mathbb{R} \rightarrow \mathbb{R}^+$ defined by

$$\zeta(0) = 1 \quad \text{and} \quad \forall y \neq 0, \quad \zeta(y) = \frac{1 - \exp(-y)}{y}, \quad (12)$$

$$\zeta'(0) = -1/2 \quad \text{and} \quad \forall y \neq 0, \quad \zeta'(y) = \frac{(1+y)\exp(-y) - 1}{y^2} = \frac{\exp(-y) - \zeta(y)}{y},$$

$$\omega(0) = 1/2 \quad \text{and} \quad \forall y \neq 0, \quad \omega(y) = \frac{\exp(-y) - 1 + y}{y^2} = \frac{1 - \zeta(y)}{y}, \quad (13)$$

$$\omega'(0) = -1/6 \quad \text{and} \quad \forall y \neq 0, \quad \omega'(y) = \frac{2(1 - \exp(-y)) - y(1 + \exp(-y))}{y^3} = \frac{2\zeta(y) - 1 - \exp(-y)}{y^2}.$$

Both functions non-increasing, diverge to $+\infty$ at negative infinity and vanish at positive infinity. Let us now enounce the main theorem for the optimal execution problem.

Theorem 4.1. *Let $\epsilon \in [0, 1)$. The optimal strategy X^* that minimizes the expected cost $\mathbb{E}[C(X)]$ among admissible strategies that liquidate x_0 assets is explicit. It is a linear combination of $(x_0, D_0, \delta_0, I, N)$ and can be written as*

$$X^* = X^{\text{OW}} + X^{\text{trend}} + X^{\text{dyn}},$$

where

- X^{OW} is the optimal strategy in the Obizhaeva and Wang [32] model, given by (7) in Theorem 2.1,
- X^{trend} is the “trend strategy”, given by (19).
- X^{dyn} is the “dynamic strategy”, given by (20).

The strategy X^{OW} is a linear function of x_0 , X^{trend} is a linear function of (D_0, δ_0) while X^{dyn} is a linear function of the processes I and N . The discontinuity times of X^{dyn} are those of N , and if N jumps at time $\tau \in (0, T)$, we have

$$(1 - \epsilon) \Delta X_{\tau}^{\text{dyn}} = \frac{1 + \rho(T - \tau)}{2 + \rho(T - \tau)} \left\{ \frac{m_1}{\rho} \Delta I_{\tau} - (1 - \nu) \Delta N_{\tau} \right\} + \frac{m_1}{2\rho} (\nu\rho - \eta) \frac{\rho(T - \tau)^2 \times \omega(\eta(T - \tau))}{2 + \rho(T - \tau)} \Delta I_{\tau}. \quad (14)$$

All explicit formulas are given in Appendix A. The value function of the problem is given by

$$\begin{aligned}
q \times \mathcal{C}(t, x, d, z, \delta, \Sigma) &= -q(z + d)x + \left[\frac{1 - \epsilon}{2 + \rho(T - t)} + \frac{\epsilon}{2} \right] x^2 + \frac{\rho(T - t)}{2 + \rho(T - t)} \left[qd - \mathcal{G}_\eta(T - t) \frac{\delta m_1}{\rho} \right] x \\
&\quad - \frac{1}{1 - \epsilon} \times \frac{\rho(T - t)/2}{2 + \rho(T - t)} \left[qd - \mathcal{G}_\eta(T - t) \frac{\delta m_1}{\rho} \right]^2 + \hat{c}_\eta(T - t) \left(\frac{\delta m_1}{\rho} \right)^2 \\
&\quad + e(T - t) \Sigma + g(T - t),
\end{aligned}$$

where for $u \in [0, T]$,

$$\begin{aligned}
\mathcal{G}_\eta(u) &= \zeta(\eta u) + \nu \rho u \omega(\eta u), \\
\hat{c}_\eta(u) &= \frac{1}{1 - \epsilon} \times (\eta - \nu \rho)^2 \frac{\rho u^3}{8} \omega'(\eta u) \zeta(\eta u).
\end{aligned}$$

The functions e and g are the unique solution of the differential equations (32) and (33) with $e(0) = g(0) = 0$.

The proof of this theorem is given in Appendix B. Let us mention here that the functions e and g admit explicit forms by the mean of the exponential integral function, that are very cumbersome. They can be obtained by using a formal calculus software such as Mathematica. Since they do not play any role to determine the optimal strategy and require several pages to be displayed, we do not give these explicit formulas. Note that they are simpler in the case $\eta = 0$, for which the explicit formulas are given by Equations (38) and (39).

The optimal strategy X^* is illustrated on Figure 1 for two different sets of parameters. It is worth to notice that the strategy is linear with respect to x_0 , D_0 , δ_0 , I and N . This property is due to the affine structure of the model and the quadratic costs. In particular, the reaction of the optimal strategy to the other trades does not depend on x_0 . The strategy X^{trend} is the part of the strategy which is proportional to D_0 and δ_0 and thus takes advantage of temporary price trends that are known at time 0. The strategy X^{dyn} is proportional to the processes I and N and describes the optimal reactions to observed price jumps. Last, let us stress that having an explicit formula for the optimal strategy is an important feature to use it in practice. Since the strategy reacts to each market order (or at least to those which trigger price moves), its computation time should be significantly lower than the typical duration between two of these orders.

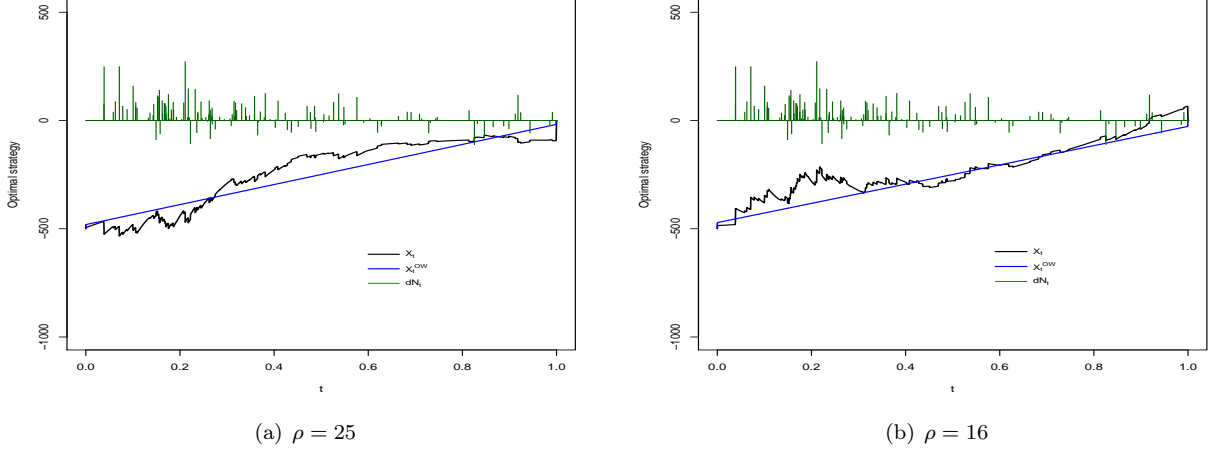


Figure 1: Optimal strategy in the Hawkes model, in black, for $q = 100$, $T = 1$, $\beta = 20$, $\iota_s = 16$, $\iota_c = 2$, $\kappa_\infty = 12$, $\epsilon = 0.3$, $\nu = 0.3$, $D_0 = 0.1$, $\kappa_0^+ = \kappa_0^- = 60$, $m_1 = 50$, $X_0 = -500$, $\mu = \text{Exp}(1/m_1)$, $\varphi_s(y) = 1.2 \times y^{0.2} + 0.5 \times y^{0.7} + 14.4 \times y$, $\varphi_c(y) = 1.2 \times y^{0.2} + 0.5 \times y^{0.7} + 0.4 \times y$ for all $y > 0$. The strategy of the Obizhaeva and Wang model is given in blue as a benchmark, and the jumps of (N_t) are plotted in green (with the same trajectory for the two graphs). On the left graph, $\iota_s < \beta < \rho$ and the strategy is based on mean-reversion: each time N jumps, X jumps in the opposite direction. On the right graph, $\rho = \iota_s < \beta$ and the strategy is trend-following.

Let us make some comments on the optimal strategy, and more precisely on how the strategic trader reacts to the orders issued by other traders. First, we observe from (20) that the block trades that immediately follow jumps of N are then compensated by the continuous trading rate. When $\varphi_s = \varphi_c$, we have $I \equiv 0$ and these block trades, as given by (14), are always opposed in sign to the market orders that they follow. For general functions φ_s and φ_c , the signs of these trades depend on the size of the last preceding jumps of N . For example, in the case where $\eta = \nu\rho$, the strategic trader makes a trade in the opposite direction if $|dN_t| > \frac{m_1}{\rho(1-\nu)}(\varphi_s - \varphi_c)(|dN_t|/m_1)$, but trades in the same way otherwise. The same conclusion holds for any parameter value when $T - t \rightarrow 0$ since $\rho(T - t)^2 \times \omega(\eta(T - t))$ vanishes. We now consider the asymptotics when the trading horizon is large: in this case, it is reasonable to assume that $\eta > 0$ which is required to get stationary intensities κ^+ and κ^- , see Section 2.3.2. Then, when $T - t \rightarrow +\infty$, the jump part of X^{dyn} given by (14) can be well approximated by

$$\frac{m_1}{2\rho} \left(1 + \frac{\nu\rho}{\eta} \right) dI_t - (1 - \nu) dN_t.$$

Therefore, the strategic trader makes a trade in the opposite direction if $|dN_t| > \frac{m_1}{2\rho(1-\nu)} \left(1 + \frac{\nu\rho}{\eta} \right) (\varphi_s - \varphi_c)(|dN_t|/m_1)$ and trades in the same direction otherwise. In the case $\iota_c = 0$ and $\varphi_s \equiv \varphi_c$ where there is only volume-independent self-excitation, we can interpret this behavior as follows: if a market buy order is relatively small, it may be a part of a big split order, and thus be followed by other buy orders that will make the price go up, and the strategic trader has interest to follow this trend. However, if a market buy order is relatively big, the price resilience effect is likely to dominate and the strategic trader has interest to trade in the opposite way.

Last, it is interesting to notice that the optimal strategy only depends on (φ_s, φ_c) through $\varphi_s - \varphi_c$. This key self-excitation function tunes the way that the strategic trader should react to other market orders.

Remark 4.1. *The MIH model with $\eta = 0$ includes the particular case of independent Poisson processes when*

$\beta = 0$ and $\varphi_s = \varphi_c \equiv 0$. In that case, if N jumps at time $\tau \in (0, T)$, we get from (14)

$$(1 - \epsilon)\Delta X_\tau^{\text{dyn}} = -\frac{1 + \rho(T - \tau)}{2 + \rho(T - \tau)} \times (1 - \nu) \Delta N_\tau.$$

Since the self-excitation effect is removed, the price is a mean-reverting process when the strategic trader is passive. Thus, each time a market order is observed, the optimal strategy consists in posting immediately a market order in the opposite direction, to arbitrage the resilience of the price. Such an obvious Price Manipulation Strategy is unrealistic, therefore modeling the order flow with Poisson processes is not satisfactory. We refer to Section 5.2 for more details.

Remark 4.2. Following Remark 2.3, a natural question is to look at the quantity x_0 that minimizes $\mathbb{E}[C(X)] + P_0 \times x_0$, i.e. the expected liquidation cost with respect to the mark-to-market value. From Theorem 4.1 we obtain easily that, at time 0, this quantity is minimal for

$$x_0^* = \frac{\rho T [qD_0 - \mathcal{G}_\eta(T) \frac{\delta_0 m_1}{\rho}]}{2(1 + \frac{\epsilon}{2}\rho T)}.$$

We can give a simple heuristic for the sign of x_0^* : when $D_0 \geq 0$ and $\delta_0 \leq 0$ the price trend is negative and it is more favorable to sell ($x_0 \geq 0$) since \mathcal{G}_η is nonnegative.

5 Price Manipulation Strategies in the MIH model

In this section, we study Price Manipulation Strategies (PMS), as introduced by Definition 2.2, in the context of the MIH model. As a matter of fact, the value function given in Theorem 4.1 can be negative even for $x_0 = 0$, which would constitute a PMS. We first determine necessary and sufficient conditions on the parameters of the model to exclude such strategies. Then, we study the particular case of Poisson processes, which may seem natural to model the order flow but allow for robust arbitrages to arise in this framework.

5.1 The Mixed-market-Impact Hawkes Martingale (MIHM) model

Theorem 2.1 gives a necessary and sufficient condition on N to exclude Price Manipulation Strategies. Here, we apply this result to identify which parameters in the Hawkes model exclude PMS. We recall the notation

$$\alpha = \iota_s - \iota_c = \int_{\mathbb{R}^+} (\varphi_s - \varphi_c)(v/m_1) \mu(dv),$$

and define the (normalized) support of μ

$$\mathcal{S}(\mu) = \{y \geq 0 \text{ s.t. } \forall \varepsilon > 0, \mu((m_1 \times y - \varepsilon, m_1 \times y + \varepsilon)) > 0\}.$$

Proposition 5.1. *The MIH model does not admit PMS if, and only if the following conditions hold*

$$\beta = \rho, \quad \alpha = (1 - \nu)\rho, \quad \varphi_s(x) - \varphi_c(x) = \alpha x \text{ for } x \in \mathcal{S}(\mu) \text{ (i.e. } m_1 I = \alpha N), \text{ and } qD_0 = \frac{m_1}{\rho} \delta_0, \quad (15)$$

or $\mu = \text{Dirac}(0)$ with $D_0 = 0$. In both cases, the optimal execution strategy is given by (7).

Note that in the case $\mu = \text{Dirac}(m_1)$ where all the jumps have the same size, one has $\mathcal{S}(\mu) = \{1\}$ thus $\varphi_s - \varphi_c$ is necessarily linear on $\mathcal{S}(\mu)$ and $\Delta I_t = \alpha \text{sgn}(\Delta N_t)$. If moreover $m_1 = 0$, we have $N \equiv 0$ and the MIH model does not depend any longer on the parameters α and β , that can then be fixed arbitrarily.

Proof. From Theorem 2.1, PMS are excluded if, and only if the price P is a martingale when $X \equiv 0$. In this case, we have from (1), (2), (3) and (10)

$$dP_t = -\rho D_t dt + \frac{1}{q} dN_t = \frac{1}{q} (dN_t - \delta_t m_1 dt) + \left(\frac{m_1}{q} \delta_t - \rho D_t \right) dt.$$

Therefore, P is a martingale if, and only if $\frac{m_1}{\rho} \delta_t = q D_t$ \mathbb{P} -a.s., dt a.e. This condition is equivalent to $q D_0 = \frac{m_1}{\rho} \delta_0$ and $q dD_t = \frac{m_1}{\rho} d\delta_t$. From (3) and (10), the latter condition is equivalent to

$$\rho q D_t = \frac{m_1}{\rho} \beta \delta_t \text{ and } (1 - \nu) dN_t = \frac{m_1}{\rho} dI_t.$$

Using (11), the second condition is equivalent to $(1 - \nu) \rho v = m_1 (\varphi_s - \varphi_c) (v/m_1)$ for all v in the support of μ , which implies the linearity of $\varphi_s - \varphi_c$ on $\mathcal{S}(\mu)$ and leads to (15). Conversely, (15) implies $\frac{m_1}{\rho} \delta_t = q D_t$, and P is then a martingale. \square

Remark 5.1. When $\beta = \rho$, $\alpha = (1 - \nu)\rho$, and $\varphi_s - \varphi_c$ is linear on $\mathcal{S}(\mu)$, we get from the previous calculations that $d(\frac{m_1}{q} \delta_t - \rho D_t) = -\rho(\frac{m_1}{q} \delta_t - \rho D_t) dt$, and therefore $\frac{m_1}{q} \delta_t - \rho D_t$ converges exponentially to zero. The condition $q D_0 = \frac{m_1}{\rho} \delta_0$ simply means that the model starts from this steady state.

One can also check directly that the optimal strategy and its cost given by Theorem 4.1 coincide with those of Theorem 2.1 when (15) holds. For clear reasons, we call Mixed-Impact Hawkes Martingale (MIHM) model the MIH model if these conditions are satisfied. Proposition 5.1 is very interesting since it makes connections between the model parameters of the MIH model in a perfect market without PMS. First, the condition $\beta = \rho$ means that the mean-reverting action of liquidity providers compensates the autocorrelation in the signs of the trades of liquidity takers; we thus reach a conclusion similar to Bouchaud et al. [11]. The condition $\alpha = (1 - \nu)\beta$ gives a link between the Hawkes kernel and the proportion $1 - \nu$ of transient price impact. When $\iota_c = 0$, α/β represents the average number of child orders coming from one market order, and is thus equal to the proportion of endogenous orders (i.e. triggered by other orders) in the market. What we obtain here is that this ratio should be equal to $1 - \nu$, which is a *a priori* different measure of endogeneity, since it gives the proportion of market impact that does not influence the low-frequency price (see Section 2.3.2). The positivity of α reflects the fact that the parameter ι_c tuning opportunistic trading should be small to avoid market instability. It is interesting to notice that if (15) holds, the stationarity condition $\iota_s + \iota_c < \beta$ derived in Section 2.3.2 is equivalent to $2\iota_c < \nu\rho$, which can be seen as a reasonable upper bound for ι_c . Last, we see that $\varphi_s - \varphi_c$ should be linear. Let us recall that φ_s and φ_c encode the dependence of the self-excitation (resp. the cross-excitation) effect on the volumes of incoming market orders. Condition (15) implies that they should have roughly the same functional form, except for a linear part which should be stronger for φ_s . However, we remind here that these conclusions are obtained in the MIH model and should be confronted to market data. We leave this empirical investigation for further research.

Of course, in practice, it would be miraculous if the calibration of the MIH model on real financial data led to parameters satisfying exactly (15). One may rather expect these parameters to be close but not exactly equal to those of the MIHM model, for the following reasons. First, there is no guarantee that fitting a model to a market with no PMS leads to a model with no PMS. Second, the MIH model ignores market frictions such as the bid-ask spread and gives some advantages to the strategic trader such as the possibility to post orders immediately after the other ones (see Stoikov and Waeber [36] for a study on the latency to execute an order). These facts make the existence of PMS more likely in the model than in reality. Third, we know that in practice, temporary arbitrage may exist at high frequencies. Therefore, there is no reason that fitted parameters follow exactly the MIHM condition (15). This justifies the potential practical usefulness of the strategy given by Theorem 4.1 to reduce execution costs when the estimated parameters deviate from the MIHM model. Let us note that Figure 1 illustrates such a case: all the parameters satisfy (15) but ρ (which should be equal to $\beta = 20$). The estimation of the MIH model on market data is left for future research.

The framework of the MIH model also gives some interesting insights for the characterization of the existence of short-time arbitrages. Let us introduce the following definition.

Definition 5.1. We say that a market admits weak Price Manipulation Strategies (wPMS) if the cost of a liquidation strategy can be reduced by posting a block trade as an immediate response to a market order issued by another trader.

Corollary 5.1. In the MIH model, the market does not admit wPMS if, and only if,

$$\beta = \rho, \alpha = (1 - \nu)\rho \text{ and } \varphi_s(x) - \varphi_c(x) = \alpha x \text{ for } x \in \mathcal{S}(\mu) \quad (16)$$

or $\mu = \text{Dirac}(0)$.

Proof. The proof is quite straightforward from Theorem 4.1. The case $\mu = \text{Dirac}(0)$ is trivial and we consider $m_1 > 0$. The jump term of the strategy (14) should be equal to zero for any $\tau \in [0, T]$. By taking $\tau = T$, we get that $\varphi_s(x) - \varphi_c(x) = (1 - \nu)\rho x$ for $x \in \mathcal{S}(\mu)$. Integrating this identity with respect to μ leads to $\alpha = \iota_s - \iota_c = (1 - \nu)\rho$. Then, from (14), we should have $(\nu\rho - \eta) \times \rho u^2 \times \omega(\eta u) = 0$ for $u \in [0, T]$ which implies $\nu\rho = \eta$. Since $\eta = \beta - \alpha = \beta - (1 - \nu)\rho$, we get $\beta = \rho$. The converse implication is obvious. \square

By Remark 5.1, the condition $qD_0 = \frac{m_1}{\rho}\delta_0$ means that the model has reached its equilibrium, which is basically the case after some time. Therefore, the conditions that exclude wPMS and PMS in the MIH model are quite the same. This is an interesting link between two different point of views. The condition “no PMS” means that there is no free source of income. The condition “no wPMS” rather brings on market stability, since it excludes trading volume coming from the response to other trades. Corollary 5.1 is a mathematical formulation of this link in our specific model.

5.2 The Poisson model

Poisson processes are often used to model the arrival of the customers in queuing theory. It is therefore natural to use them to model the flow of market orders, as it has been made for example by Bayraktar and Ludkovski [9] or Cont and de Larrard [13] in different frameworks.

Here, in the Poisson model, N^+ and N^- are two i.i.d. independent compound Poisson processes of respective constant jump rates κ_0^+ and κ_0^- , with the same jump law μ . It is a particular case of the MIH model when $\beta = 0$, $\varphi_s \equiv 0$ and $\varphi_c \equiv 0$, which implies $\eta = 0$. Thus, the optimal strategy and value function in this case can be deduced from Theorem 4.1 (see also Remark 4.1).

First, let us note that the Poisson model cannot satisfy the condition (15), except in the case $\rho = 0$, where there is only permanent price impact, which is not relevant in this context. Thus, we know *a priori* that PMS are possible. However, we specify in what follows that a Poisson order flow creates very simple and robust arbitrages. First, we put aside the case $\kappa_0^+ \neq \kappa_0^-$ where the trend on the price leads to obvious PMS, and consider now the more interesting case $\kappa_0^+ = \kappa_0^-$, and we simply denote by κ_0 the common intensity.

A natural choice to get a PMS is of course to consider the optimal strategy given by Theorem 4.1 when liquidating $x_0 = 0$ assets. A remarkable feature of this optimal strategy in the Poisson case is that it only depends on the process N , and does not depend directly on the law of the jumps and their intensity. Then, when applying the optimal strategy, mainly two quantities have to be known: qD_0 and ρ . We denote by $\mathcal{C}_0(D_0)$ the cost of the optimal strategy and obtain from Theorem 4.1 in this case:

$$(1 - \epsilon)q \times \mathcal{C}_0(D_0) = -\frac{\rho T/2}{2 + \rho T} q^2 D_0^2 - (1 - \nu)^2 2\kappa_0 m_2 \left[\frac{T}{2} - \frac{1}{\rho} \ln \left(1 + \frac{\rho T}{2} \right) \right]. \quad (17)$$

In fact, PMS are very robust in this framework. The following proposition shows that even if qD_0 and ρ are unknown, one can construct a such a strategy. This indicates that in our framework with a linear price impact and an exponential resilience, compound Poisson processes are not suitable to model the order flow.

Proposition 5.2. Let $\kappa_0^+ = \kappa_0^- = \kappa_0 > 0$ and $\lambda \in (0, 1)$. The following round-trip strategy $X_0^\lambda = X_{T^+}^\lambda = 0$ defined by

$$X_{\tau^+}^\lambda - X_\tau^\lambda = -\frac{1-\nu}{1-\epsilon} \times \lambda(N_\tau - N_{\tau-})$$

at each jump of N is a PMS. Its average cost is given by

$$\mathbb{E}[C(X^\lambda)] = 2\lambda(1-\lambda) \frac{\kappa_0 m_2 (1-\nu)^2}{q(1-\epsilon)} \left[\frac{1 - \exp(-\rho T)}{\rho} - T \right] < 0,$$

and the best choice is to take $\lambda = 1/2$.

Proof. From Remark 2.7, it is sufficient to focus on the case $\nu = \epsilon = 0$ and $q = 1$. In this case, we have

$$C(X) = \int_{[0, T)} D_u dX_u^\lambda + \frac{1}{2} \sum_{0 \leq \tau < T} (\Delta X_\tau^\lambda)^2 - D_T X_T^\lambda + \frac{1}{2} (X_T^\lambda)^2,$$

with $D_t = D_0 + \int_0^t \exp(-\rho(t-s)) dN_s + \int_0^t \exp(-\rho(t-s)) dX_s^\lambda$. From $\int_{[0, T)} D_u dX_u^\lambda = -\lambda \sum_{0 \leq \tau < T} [D_\tau - \Delta N_\tau + (\Delta N_\tau)^2]$, we get $\mathbb{E}[\int_{[0, T)} D_u dX_u^\lambda] = -\lambda \mathbb{E}[\sum_{0 \leq \tau < T} (\Delta N_\tau)^2] = -2\lambda \kappa_0 m_2 T$. Since $X_T^\lambda = -\lambda N_T$ and $D_T = D_0 + (1 - \lambda) \int_0^T \exp(-\rho(T-s)) dN_s$ a.s., we have $\mathbb{E}[(X_T^\lambda)^2] = 2\lambda^2 \kappa_0 m_2 T$ and

$$\mathbb{E}[-D_T X_T^\lambda] = \lambda(1-\lambda) \mathbb{E} \left[N_T \int_0^T \exp(-\rho(T-s)) dN_s \right] = 2\lambda(1-\lambda) \kappa_0 m_2 \frac{1 - \exp(-\rho T)}{\rho}.$$

This eventually yields

$$\begin{aligned} \mathbb{E}[C(X^\lambda)] &= -2\lambda \kappa_0 m_2 T + \lambda^2 \kappa_0 m_2 T + 2\lambda(1-\lambda) \kappa_0 m_2 \frac{1 - \exp(-\rho T)}{\rho} + \lambda^2 \kappa_0 m_2 T \\ &= 2\lambda(1-\lambda) \kappa_0 m_2 \left(\frac{1 - \exp(-\rho T)}{\rho} - T \right). \end{aligned}$$

□

A Explicit formulas for the optimal strategy

We use the function

$$L(r, \lambda, t) := r \int_0^t \frac{\exp(\lambda s)}{2 + rs} ds = \exp(-2\lambda/r) \left[\mathcal{E} \left(\frac{\lambda}{r} (2 + rt) \right) - \mathcal{E} \left(\frac{2\lambda}{r} \right) \right], \quad (18)$$

where $\mathcal{E}(y) = -\int_{-y}^{+\infty} \frac{e^{-u}}{u} du$ is the exponential integral of y , in terms of Cauchy principal value if $y > 0$. Since we only consider differences $\mathcal{E}(y) - \mathcal{E}(y')$ with either $y, y' > 0$ or $y, y' < 0$, we will only consider proper integrals. The function \mathcal{E} is standard and is implemented in many packages such as the Boost C++ library. Thus, L can be evaluated as a closed formula.

We refer to (12) and (13) for the definitions of ζ and ω .

Auxiliary functions: For $0 \leq s \leq t \leq T$,

$$\begin{aligned} \phi_\eta(t) &= \frac{1}{2(2 + \rho(T-t))} \times \left[1 + \exp(-\eta(T-t)) + \nu \rho(T-t) \zeta(\eta(T-t)) \right. \\ &\quad \left. + \frac{\beta}{\rho} [2 + \rho(T-t) \times \{1 + \zeta(\eta(T-t)) + \nu \rho(T-t) \omega(\eta(T-t))\}] \right], \end{aligned}$$

$$\begin{aligned}
\Phi_0(s, t) &= \left[\frac{\beta}{\rho} + \frac{\nu}{2} \left(\frac{1}{2} - \frac{\beta}{\rho} \right) \right] \times \frac{\exp(-\beta s) - \exp(-\beta t)}{\beta} \\
&\quad + (1 - \nu) \left(1 - \frac{\beta}{\rho} \right) \times \frac{\exp(-\beta T)}{\rho} \times [L(\rho, \beta, T - s) - L(\rho, \beta, T - t)] \\
&\quad + \frac{\nu}{4} [(T - s) \exp(-\beta s) - (T - t) \exp(-\beta t)],
\end{aligned}$$

and for $\eta \neq 0$,

$$\begin{aligned}
\Phi_\eta(s, t) &= \frac{1}{2} \left(\frac{1}{\rho} + \frac{\nu}{\eta} \right) \times [\exp(-\beta s) - \exp(-\beta t)] \\
&\quad + \frac{\exp(-\beta T)}{2\rho} \times \left[1 + \frac{\nu(\rho - 2\beta)}{\eta} + \frac{\beta}{\eta} \left(1 - \frac{\nu\rho}{\eta} \right) \right] \times [L(\rho, \beta, T - s) - L(\rho, \beta, T - t)] \\
&\quad + \frac{\exp(-\beta T)}{2\rho} \times \left[1 - \frac{\nu\rho}{\eta} - \frac{\beta}{\eta} \left(1 - \frac{\nu\rho}{\eta} \right) \right] \times [L(\rho, \alpha, T - s) - L(\rho, \alpha, T - t)].
\end{aligned}$$

We now give the explicit formulas for the whole optimal strategy. They are valid for all $\eta \in \mathbb{R}$.

Trend strategy:

$$\begin{aligned}
(1 - \epsilon)\Delta X_0^{\text{trend}} &= \frac{\frac{\delta_0 m_1}{2\rho} \times [2 + \rho T \times \{1 + \zeta(\eta T) + \nu\rho T \omega(\eta T)\}] - [1 + \rho T]qD_0}{2 + \rho T}, \\
(1 - \epsilon)\Delta X_T^{\text{trend}} &= \frac{\delta_0 m_1}{2\rho} \times \left[\frac{2 + \rho T \times \{1 + \zeta(\eta T) + \nu\rho T \omega(\eta T)\}}{2 + \rho T} - 2\rho \Phi_\eta(0, T) - 2\exp(-\beta T) \right] \\
&\quad + \frac{qD_0}{2 + \rho T}, \tag{19}
\end{aligned}$$

and, on $(0, T)$,

$$\begin{aligned}
(1 - \epsilon)dX_t^{\text{trend}} &= \frac{\delta_0 m_1}{2\rho} \times \left[\frac{2 + \rho T \times \{1 + \zeta(\eta T) + \nu\rho T \omega(\eta T)\}}{2 + \rho T} - 2\rho \Phi_\eta(0, t) - 2\phi_\eta(t) \exp(-\beta t) \right] \rho dt \\
&\quad + \frac{qD_0}{2 + \rho T} \rho dt.
\end{aligned}$$

Dynamic strategy:

$$\begin{aligned}
(1 - \epsilon)\Delta X_0^{\text{dyn}} &= 0, \\
(1 - \epsilon)\Delta X_T^{\text{dyn}} &= -m_1 \left[\Theta_{\chi_T} \Phi_\eta(\tau_{\chi_T}, T) + \sum_{i=1}^{\chi_T-1} \Theta_i \Phi_\eta(\tau_i, \tau_{i+1}) \right] + \sum_{0 < \tau \leq T} \frac{(1 - \nu) \Delta N_\tau}{2 + \rho(T - \tau)} \\
&\quad + \frac{m_1}{2\rho} \times \sum_{0 < \tau \leq T} \frac{2 + \rho(T - \tau) \times \{1 + \zeta(\eta(T - \tau)) + \nu\rho(T - \tau) \omega(\eta(T - \tau))\}}{2 + \rho(T - \tau)} \Delta I_\tau \\
&\quad - \frac{m_1}{\rho} \Theta_{\chi_T} \exp(-\beta T), \tag{20}
\end{aligned}$$

and, on $(0, T)$,

$$\begin{aligned}
(1 - \epsilon)dX_t^{\text{dyn}} &= -m_1 \phi_\eta(t) \Theta_{\chi_t} \exp(-\beta t) dt + \left[\sum_{0 < \tau \leq t} \frac{(1 - \nu)\Delta N_\tau}{2 + \rho(T - \tau)} \right] \rho dt \\
&+ \left[\sum_{0 < \tau \leq t} \frac{2 + \rho(T - \tau) \times \{1 + \zeta(\eta(T - \tau)) + \nu\rho(T - \tau) \omega(\eta(T - \tau))\}}{2 + \rho(T - \tau)} \Delta I_\tau \right] \frac{m_1}{2} dt \\
&- \left[\Theta_{\chi_t} \Phi_\eta(\tau_{\chi_t}, t) + \sum_{i=1}^{\chi_t-1} \Theta_i \Phi_\eta(\tau_i, \tau_{i+1}) \right] \rho m_1 dt \\
&+ \frac{1 + \rho(T - t)}{2 + \rho(T - t)} \left\{ \frac{m_1}{\rho} dI_t - (1 - \nu) dN_t \right\} + \frac{m_1}{2\rho} (\nu\rho - \eta) \times \frac{\rho(T - t)^2 \times \omega(\eta(T - t))}{2 + \rho(T - t)} dI_t.
\end{aligned}$$

B Proof for the optimal control problem (results of Theorem 4.1 and Appendix A)

B.1 Notations and methodology

The jump intensity of the process (N_t) is characterized by the càdlàg Markovian process (δ_t, Σ_t) defined by (10), taking values in $\mathbb{R} \times \mathbb{R}^+$. The state variable of the problem is then $(X_t, D_t, S_t, \delta_t, \Sigma_t)$, and the control is $X_t - x_0$, i.e. the variation of the position of the strategic trader, $(X_t)_{t \in [0, T]}$ being an admissible strategy as described in Definition 2.1. The control program is thus to minimize $\mathbb{E}[C(0, X)]$ over all admissible strategies, where the cost $C(t, X)$ of the strategy X between t and T is given by

$$C(t, X) = \int_{[t, T)} P_u dX_u + \frac{1}{2q} \sum_{t \leq \tau < T} (\Delta X_\tau)^2 - P_T X_T + \frac{1}{2q} X_T^2.$$

The final value at time $t = T$ is the cost of a market order of signed volume $\Delta X_T = -X_T$ (so that $X_{T+} = X_T + \Delta X_T = 0$). At time t , the price P_t depends on D_t and S_t which in turn depend on $(X_u)_{u \in [0, t]}$. Let us define \mathcal{A}_t the set of admissible strategies on $[t, T]$, with $t \in [0, T]$. The value function of the problem is

$$\mathcal{C}(t, x, d, z, \delta, \Sigma) = \inf_{X \in \mathcal{A}_t} \mathbb{E}[C(t, X)]$$

with $X_t = x$, $D_t = d$, $S_t = z$, $\delta_t = \delta$ and $\Sigma_t = \Sigma$. In order to determine analytically the value function and the optimal control of the problem, we use the probabilistic formulation of the verification theorem. We determine *a priori* a continuously differentiable function $\mathcal{C}(t, x, d, z, \delta, \Sigma)$ and an admissible strategy X^* and then we verify that

$$\Pi_t(X) := \int_0^t P_u dX_u + \frac{1}{2q} \sum_{0 \leq \tau < t} (\Delta X_\tau)^2 + \mathcal{C}(t, X_t, D_t, S_t, \delta_t, \Sigma_t) \quad (21)$$

is a submartingale for any admissible strategy X , and that $\Pi_t(X^*)$ is a martingale. We proceed in three steps:

1. We define a suitable function \mathcal{C} , and derive a set of ODEs on its coefficients which is a necessary condition for \mathcal{C} to be the value function of the problem.
2. We solve the set of ODEs.

3. Using the results of the previous steps, we derive the strategy X^* such that $\Pi_t(X^*)$ is a martingale.

The verification argument then yields that $\mathcal{C}(t, x, d, z, \delta, \Sigma)$ is the value function and that X^* is optimal. Without loss of generality, we can assume that $q = 1$ by using Remark 2.4.

B.2 Necessary conditions on the value function

We search a cost function \mathcal{C} as a generic quadratic form of the variables x, d, z, δ, Σ with time-dependent coefficient (the variable z symbolizes the current value of the fundamental price S_t). As we see further, we need \mathcal{C} to verify $\partial_x \mathcal{C} + (1 - \epsilon) \partial_d \mathcal{C} + \epsilon \partial_z \mathcal{C} + d + z = 0$: it is thus necessary that \mathcal{C} is a quadratic form of $(d - (1 - \epsilon)x)$, $(z - \epsilon x)$, δ and Σ , plus a term $-(d + z)^2/2$. We define

$$\begin{aligned} \mathcal{C}(t, x, d, z, \delta, \Sigma) &= a(T-t)(d - (1 - \epsilon)x)^2 + \frac{1}{2}(z - \epsilon x)^2 + (d - (1 - \epsilon)x)(z - \epsilon x) - \frac{(d + z)^2}{2} \\ &\quad + b(T-t) \delta (d - (1 - \epsilon)x) + c(T-t) \delta^2 + e(T-t) \Sigma + g(T-t), \end{aligned} \quad (22)$$

with $a, b, c, e, g: \mathbb{R}^+ \rightarrow \mathbb{R}$ continuously differentiable functions. We choose the limit condition $\mathcal{C}(T, x, d, z, \delta, \Sigma) = -(d + z)x + x^2/2 = \frac{1}{2}(d + z - x)^2 - (d + z)^2/2$, which is the cost of a trade of signed volume $-x$. We thus have

$$a(0) = \frac{1}{2}, \quad b(0) = c(0) = e(0) = g(0) = 0.$$

Let us note that other terms should be added in equation (22) for \mathcal{C} to be a generic quadratic form. The five terms

$$h_1(T-t) (d - (1 - \epsilon)x) + h_2(T-t) \Sigma(d - (1 - \epsilon)x) + h_3(T-t) \delta \Sigma + h_4(T-t) \delta + h_5(T-t)(z - \epsilon x)$$

have to be equal to zero since $\mathcal{C}(t, x, d, z, \delta, \Sigma) = \mathcal{C}(t, -x, -d, -z, -\delta, \Sigma)$ by using Remark 2.4 and the fact that the buy and sell orders play a symmetric role. For the term in Σ^2 , we checked in prior calculations that it is necessarily associated to a zero coefficient. For $\Delta x \in \mathbb{R}$, we have

$$\mathcal{C}(t, x + \Delta x, d + (1 - \epsilon)\Delta x, z + \epsilon\Delta x, \delta, \Sigma) - \mathcal{C}(t, x, d, z, \delta, \Sigma) = -(d + z) \times \Delta x - \frac{(\Delta x)^2}{2}. \quad (23)$$

In what follows, we drop the dependence of $\mathcal{C}(t, X_t, D_t, S_t, \delta_t, \Sigma_t)$ on $(t, X_t, D_t, S_t, \delta_t, \Sigma_t)$ to obtain less cumbersome expressions. The process $\mathcal{C}(t, X_t, D_t, S_t, \delta_t, \Sigma_t)$ is $\text{l\`a}d\text{l}\grave{a}g$, and with the notations of Remark 2.5, we have by using (23)

$$\begin{aligned} d\mathcal{C} &= \partial_t \mathcal{C} dt + \partial_x \mathcal{C} dX_t^c + \partial_d \mathcal{C} \left(-\rho D_t dt + (1 - \epsilon) dX_t^c \right) + \partial_z \mathcal{C} \epsilon dX_t^c \\ &\quad - \beta \delta_t \partial_\delta \mathcal{C} dt - \beta (\Sigma_t - 2\kappa_\infty) \partial_\Sigma \mathcal{C} dt \\ &\quad + \left[\mathcal{C}(t, X_t, D_{t-} + (1 - \nu)\Delta N_t, S_{t-} + \nu\Delta N_t, \delta_{t-} + \Delta I_t, \Sigma_{t-} + \Delta \bar{I}_t) - \mathcal{C}(t, X_t, D_{t-}, S_{t-}, \delta_{t-}, \Sigma_{t-}) \right] \\ &\quad - (D_t + S_t) \Delta X_t - \frac{(\Delta X_t)^2}{2}. \end{aligned}$$

where we refer to (11) for the definitions of I and \bar{I} . The definition of $\Pi(X)$ given by (21) yields $d\Pi_t(X) = (D_t + S_t)dX_t^c + (D_t + S_t)\Delta X_t + (\Delta X_t)^2/2 + d\mathcal{C}$. We define the continuous finite variation process $(A_t^X)_{t \in (0, T)}$ such that $A_{0+}^X = \mathcal{C}(0, X_{0+}, D_{0+}, S_{0+}, \delta_0, \Sigma_0)$ and for $t \in (0, T)$

$$\begin{aligned} dA_t^X &= (D_t + S_t) dX_t^c + Z(t, X_t, D_t, S_t, \delta_t, \Sigma_t) dt \\ &\quad + \partial_t \mathcal{C} dt + \partial_x \mathcal{C} dX_t^c + \partial_d \mathcal{C} \left(-\rho D_t dt + (1 - \epsilon) dX_t^c \right) + \partial_z \mathcal{C} \epsilon dX_t^c \\ &\quad - \beta \delta_t \partial_\delta \mathcal{C} dt - \beta (\Sigma_t - 2\kappa_\infty) \partial_\Sigma \mathcal{C} dt, \end{aligned}$$

where, for $V \sim \mu$,

$$Z(t, x, d, z, \delta, \Sigma) :=$$

$$\begin{aligned} & \frac{\Sigma + \delta}{2} \times \mathbb{E}[\mathcal{C}(t, x, d + (1 - \nu)V, z + \nu V, \delta + (\varphi_s - \varphi_c)(V/m_1), \Sigma + (\varphi_s + \varphi_c)(V/m_1)) - \mathcal{C}(t, x, d, z, \delta, \Sigma)] \\ & + \frac{\Sigma - \delta}{2} \times \mathbb{E}[\mathcal{C}(t, x, d - (1 - \nu)V, z - \nu V, \delta - (\varphi_s - \varphi_c)(V/m_1), \Sigma + (\varphi_s + \varphi_c)(V/m_1)) - \mathcal{C}(t, x, d, z, \delta, \Sigma)]. \end{aligned}$$

Then, $\Pi(X) - A^X$ is a martingale (let us note that almost surely, dt -a.e. on $(0, T)$, $Z(t, X_t, D_t, S_t, \delta_t, \Sigma_t) = Z(t, X_t, D_t, S_t, \delta_t, \Sigma_t)$). This yields that $\Pi(X)$ is a submartingale (resp. a martingale) iff A^X is increasing (resp. constant). From (23), we obtain $\partial_x \mathcal{C}(t, x, d, z, \delta, \Sigma) + (1 - \epsilon) \partial_d \mathcal{C}(t, x, d, z, \delta, \Sigma) + \epsilon \partial_z \mathcal{C}(t, x, d, z, \delta, \Sigma) + d + z = 0$, and then

$$dA_t^X = \left\{ \partial_t \mathcal{C} - \rho D_t \partial_d \mathcal{C} + Z(t, X_t, D_t, S_t, \delta_t, \Sigma_t) - \beta \delta_t \partial_\delta \mathcal{C} - \beta (\Sigma_t - 2\kappa_\infty) \partial_\Sigma \mathcal{C} \right\} dt. \quad (24)$$

Given the quadratic nature of the problem, we search a process A^X of the form

$$dA_t^X = \frac{\rho}{1 - \epsilon} dt \times \left[j(T - t)(D_t - (1 - \epsilon)X_t) - D_t + k(T - t) \delta_t \right]^2, \quad (25)$$

with $j, k : \mathbb{R}^+ \rightarrow \mathbb{R}$ continuously differentiable functions, in order to obtain an non-decreasing process A^X that can be constant for a specific strategy X^* . Let us note $Y_t := D_t - (1 - \epsilon)X_t$, $\Xi_t := S_t - \epsilon X_t$, $y := d - (1 - \epsilon)x$, $\xi := z - \epsilon x$. Since $d + z = y + \xi + x = \xi + \frac{d - \epsilon y}{1 - \epsilon}$, we have

$$\begin{aligned} \partial_t \mathcal{C}(t, x, d, z, \delta, \Sigma) &= -\dot{a} y^2 - \dot{b} \delta y - \dot{c} \delta^2 - \dot{e} \Sigma - \dot{g}, \\ -\rho d \partial_d \mathcal{C}(t, x, d, z, \delta, \Sigma) &= -\left(2\rho a + \frac{\rho \epsilon}{1 - \epsilon} \right) dy + \frac{\rho}{1 - \epsilon} d^2 - \rho b \delta d, \\ -\beta \delta \partial_\delta \mathcal{C}(t, x, d, z, \delta, \Sigma) &= -\beta b \delta y - 2\beta c \delta^2, \\ -\beta (\Sigma - 2\kappa_\infty) \partial_\Sigma \mathcal{C}(t, x, d, z, \delta, \Sigma) &= -\beta e \Sigma + 2\beta \kappa_\infty e, \end{aligned}$$

Let $V \sim \mu$. One has

$$\mathbb{E}[(\varphi_s - \varphi_c)(V/m_1)] = \iota_s - \iota_c = \alpha \quad , \quad \mathbb{E}[(\varphi_s + \varphi_c)(V/m_1)] = \iota_s + \iota_c = \alpha + 2\iota_c.$$

Thus,

$$\begin{aligned} & \mathbb{E}[\mathcal{C}(t, x, d + (1 - \nu)V, z + \nu V, \delta + (\varphi_s - \varphi_c)(V/m_1), \Sigma + (\varphi_s + \varphi_c)(V/m_1)) - \mathcal{C}(t, x, d, z, \delta, \Sigma)] \\ &= a [(1 - \nu)^2 m_2 + 2(1 - \nu)m_1 y] + \frac{\nu^2}{2} m_2 + \nu m_1 \xi \\ & \quad + \nu(1 - \nu)m_2 + \nu m_1 y + (1 - \nu)m_1 \xi - \frac{1}{2} \left(m_2 + 2 m_1 \xi + \frac{2m_1}{1 - \epsilon} d - \frac{2\epsilon m_1}{1 - \epsilon} y \right) \\ & \quad + b [(1 - \nu)m_1 \delta + \alpha y + \tilde{\alpha}(1 - \nu)] + c [\alpha_2 + 2\alpha \delta] + (\alpha + 2\iota_c)e, \end{aligned}$$

with

$$\tilde{\alpha} = \mathbb{E}[V \times (\varphi_s - \varphi_c)(V/m_1)] \quad , \quad \alpha_2 = \mathbb{E}[(\varphi_s - \varphi_c)^2(V/m_1)]. \quad (26)$$

These quantities $\tilde{\alpha}$ and α_2 are finite by assumption. This gives

$$\begin{aligned} Z(t, x, d, z, \delta, \Sigma) &= \left(m_1 \times \left[2(1 - \nu)a + \nu + \frac{\epsilon}{1 - \epsilon} \right] + \alpha b \right) \delta y - \frac{m_1}{1 - \epsilon} \delta d \\ & \quad + [(1 - \nu)m_1 b + 2\alpha c] \delta^2 \\ & \quad + \left(m_2 \times \left[(1 - \nu)^2 a + \nu(1 - \nu/2) - \frac{1}{2} \right] + \tilde{\alpha}(1 - \nu)b + \alpha_2 c + (\alpha + 2\iota_c)e \right) \Sigma, \end{aligned}$$

where we consider \mathcal{C} as a function of the variables $t, x, d, z, \delta, \Sigma$ as in equation (24), and substitute $d - (1 - \epsilon)x$ by y and $z - \epsilon x$ by ξ in the results. We then make the change of variables $(x, d, z, \delta, \Sigma) \rightarrow (y, d, \xi, \delta, \Sigma)$, and we identify each term of equations (24) and (25):

$$\text{(Eq. } dy\text{):} \quad - \left(2\rho a + \frac{\rho\epsilon}{1-\epsilon} \right) = -\frac{2\rho}{1-\epsilon} j.$$

$$\text{(Eq. } y^2\text{):} \quad -\dot{a} = \frac{\rho}{1-\epsilon} j^2.$$

(Eq. dy) yields $j = (1 - \epsilon)a + \frac{\epsilon}{2}$. We input this relation in (Eq. y^2) and we have $\dot{j} = (1 - \epsilon)\dot{a} = -\rho j^2$ thus $j(u) = \frac{1}{2+\rho u}$ since $j(0) = (1 - \epsilon)a(0) + \frac{\epsilon}{2} = \frac{1}{2}$. This yields $a(u) = \frac{1}{1-\epsilon} \left(\frac{1}{2+\rho u} - \frac{\epsilon}{2} \right)$ with (Eq. dy).

$$\text{(Eq. } \delta y\text{):} \quad -\dot{b} - \beta b + \alpha b + m_1 \times \left[2(1 - \nu)a + \nu + \frac{\epsilon}{1-\epsilon} \right] = \frac{2\rho}{1-\epsilon} j k.$$

$$\text{(Eq. } \delta d\text{):} \quad -\rho b - \frac{m_1}{1-\epsilon} = -\frac{2\rho}{1-\epsilon} k,$$

which yields $k(u) = \frac{1-\epsilon}{2} b(u) + \frac{m_1}{2\rho}$. Plugging equation (28) in (Eq. δy), we have $\dot{b} = -(\beta - \alpha)b - \frac{2\rho}{1-\epsilon} j \left(\frac{1-\epsilon}{2} b + \frac{m_1}{2\rho} \right) + m_1 \left[2(1 - \nu)a + \nu + \frac{\epsilon}{1-\epsilon} \right]$, and since $j/(1 - \epsilon) = a + \epsilon/[2(1 - \epsilon)]$, we have

$$\dot{b}(u) = \left[-(\beta - \alpha) - \frac{\rho}{2+\rho u} \right] b(u) + \frac{m_1}{1-\epsilon} \times \frac{1+\nu\rho u}{2+\rho u}.$$

$$\text{(Eq. } \delta^2\text{):} \quad -\dot{c} - 2\beta c + 2\alpha c + (1 - \nu)m_1 b = \frac{\rho}{1-\epsilon} k^2.$$

$$\text{(Eq. } \Sigma\text{):} \quad -\dot{e} - \beta e + (\alpha + 2\iota_c)e + m_2 \times \left[(1 - \nu)^2 a + \nu(1 - \nu/2) - \frac{1}{2} \right] + \tilde{\alpha}(1 - \nu)b + \alpha_2 c = 0.$$

We have $2(1 - \epsilon) \times \left[(1 - \nu)^2 a + \nu(1 - \nu/2) - \frac{1}{2} \right] = 2(1 - \nu)^2/(2 + \rho u) - (1 - \nu)^2\epsilon + \nu(2 - \nu)(1 - \epsilon) - (1 - \epsilon)$, thus

$$\dot{e}(u) = -(\beta - \alpha - 2\iota_c)e(u) + \tilde{\alpha}(1 - \nu)b(u) + \alpha_2 c(u) + \frac{(1-\nu)^2 m_2}{1-\epsilon} \times \left[\frac{1}{2+\rho u} - \frac{1}{2} \right]$$

$$\text{(Eq. constant):} \quad -\dot{g} + 2\beta\kappa_\infty e = 0.$$

We obtain two conditions on the coefficients of the process A^X

$$j(u) = \frac{1}{2 + \rho u}, \tag{27}$$

$$k(u) = \frac{1 - \epsilon}{2} b(u) + \frac{m_1}{2\rho}, \tag{28}$$

and the following set of necessary conditions on the coefficients of \mathcal{C}

$$a(u) = \frac{1}{1 - \epsilon} \left(\frac{1}{2 + \rho u} - \frac{\epsilon}{2} \right), \tag{29}$$

$$\dot{b}(u) = \left[-(\beta - \alpha) - \frac{\rho}{2 + \rho u} \right] b(u) + \frac{m_1}{1 - \epsilon} \times \frac{1 + \nu\rho u}{2 + \rho u}, \tag{30}$$

$$\dot{c}(u) = -2(\beta - \alpha) c(u) + (1 - \nu)m_1 b(u) - \frac{\rho}{1 - \epsilon} k(u)^2, \tag{31}$$

$$\dot{e}(u) = -(\beta - \alpha - 2\iota_c)e(u) + \tilde{\alpha}(1 - \nu)b(u) + \alpha_2 c(u) + \frac{(1 - \nu)^2 m_2}{1 - \epsilon} \times \left[\frac{1}{2 + \rho u} - \frac{1}{2} \right], \tag{32}$$

$$\dot{g}(u) = 2\beta\kappa_\infty e(u), \tag{33}$$

$$b(0) = c(0) = e(0) = g(0) = 0.$$

The resolution of this set of equations determines entirely the function $\mathcal{C}(t, x, d, z, \delta, \Sigma)$ defined in (22). This is the purpose of the next step of this proof. Let us note that at this stage, we already know that the system

given by Equations (27) to (33) admits a unique solution, and that the function \mathcal{C} which solves the system is the value function of the problem by using the verification argument.

B.3 Resolution of the system of ODEs

First of all, we use Equation (29) to simplify the function \mathcal{C} . The constant term (w.r.t. the time variable t) in equation (22) is $\frac{1}{2}(z - \epsilon x)^2 + (d - (1 - \epsilon)x)(z - \epsilon x) - \frac{(d+z)^2}{2} = -zx - \frac{d^2}{2} - \epsilon dx + \left[\frac{\epsilon}{2} + \frac{\epsilon}{2}(1 - \epsilon)\right] x^2$, thus the sum of $a(T-t)(d - (1 - \epsilon)x)^2$ and this constant term can be rewritten as

$$-(z + d)x + \left[\frac{1 - \epsilon}{2 + \rho(T-t)} + \frac{\epsilon}{2} \right] x^2 - \frac{1}{1 - \epsilon} \times \frac{\rho(T-t)/2}{2 + \rho(T-t)} d^2 + \frac{\rho(T-t)}{2 + \rho(T-t)} dx. \quad (34)$$

We note $\eta = \beta - \alpha$. To solve equation (30), we search a solution of the form $b(u) = \tilde{b}(u) \times \exp(-\eta u)/(2 + \rho u)$. This yields $\dot{\tilde{b}}(u) = \frac{m_1}{1 - \epsilon} \times (1 + \nu \rho u) \times \exp(\eta u)$. Using the respective definitions (12) and (13) of the functions ζ and ω , it is easy to see that for all $\eta \in \mathbb{R}$,

$$\exp(-\eta u) \int_0^u (1 + \nu \rho s) \exp(\eta s) ds = u\zeta(\eta u) + \nu \rho u^2 \omega(\eta u).$$

Since $\tilde{b}(0) = 2b(0) = 0$, we obtain

$$b(u) = \frac{m_1 u}{1 - \epsilon} \times \frac{\zeta(\eta u) + \nu \rho u \omega(\eta u)}{2 + \rho u} = \frac{1}{1 - \epsilon} \times \frac{\rho u}{2 + \rho u} \times \frac{m_1}{\rho} \mathcal{G}_\eta(u), \quad (35)$$

where

$$\mathcal{G}_\eta(u) := \zeta(\eta u) + \nu \rho u \omega(\eta u).$$

Equation (28) then gives

$$k(u) = \frac{m_1}{2\rho} \times \frac{2 + \rho u \times \{1 + \zeta(\eta u) + \nu \rho u \omega(\eta u)\}}{2 + \rho u}. \quad (36)$$

The remaining functions c , e and g do not play any role to determine the optimal strategy, and their expressions are harder to obtain. Let us first consider the case $\eta \neq 0$. After some tedious calculations, we can show that the function c that solves (31) with $c(0) = 0$ is given by:

$$c(u) = -\frac{1}{1 - \epsilon} \times \frac{\rho u/2}{2 + \rho u} \times \frac{m_1^2}{\rho^2} \mathcal{G}_\eta(u)^2 - \frac{m_1^2}{8(1 - \epsilon)\rho} \times \left(1 - \frac{\nu \rho}{\eta}\right)^2 \times u\zeta(\eta u) \times [1 + \exp(-\eta u) - 2\zeta(\eta u)]. \quad (37)$$

For the functions e and g , we recall here that they admit explicit but very cumbersome formulas that can be obtained by using a formal calculus software. In the case $\eta = 0$, the resolution of the ODEs is easier, and we get

$$\begin{aligned} c(u) &= -\frac{(1 - \nu)^2}{1 - \epsilon} \times \frac{m_1^2}{\rho^2} \times \left[\frac{1}{2} - \frac{1}{2 + \rho u} \right] - \frac{\nu m_1^2}{\rho^2(1 - \epsilon)} \times \left[\left(\frac{1}{2} - \frac{\nu}{4} \right) \rho u + \frac{\nu}{8} \rho^2 u^2 + \frac{\nu}{48} \rho^3 u^3 \right], \\ e(u) &= -\frac{(1 - \nu)^2}{1 - \epsilon} \times \left(m_2 - \frac{m_1(2\tilde{\alpha}\rho - \alpha_2 m_1)}{\rho^2} \right) \times \left[\frac{\mathcal{I}_0(u)}{2} - \frac{\exp(2\iota_c u)}{\rho} L(\rho, -2\iota_c, u) \right] \\ &\quad + \frac{\nu(1 - \nu)m_1}{2\rho^2(1 - \epsilon)} \times \left(\tilde{\alpha} - \frac{\alpha_2 m_1}{\rho} \right) \times \rho^2 \mathcal{I}_1(u) - \frac{\alpha_2 \nu^2 m_1^2}{4\rho^3(1 - \epsilon)} \times \left[\rho^2 \mathcal{I}_1(u) + \frac{1}{2} \rho^3 \mathcal{I}_2(u) + \frac{1}{12} \rho^4 \mathcal{I}_3(u) \right], \\ g(u) &= -2\beta\kappa_\infty \times \frac{(1 - \nu)^2}{1 - \epsilon} \times \left(m_2 - \frac{m_1(2\tilde{\alpha}\rho - \alpha_2 m_1)}{\rho^2} \right) \times \left\{ \frac{\mathcal{I}_1(u)}{2} - \frac{1}{2\iota_c \rho} \times \left[\exp(2\iota_c u) L(\rho, -2\iota_c, u) - \ln \left(1 + \frac{\rho u}{2} \right) \right] \right\} \\ &\quad + \frac{\beta\kappa_\infty \nu(1 - \nu)m_1}{2\rho^3(1 - \epsilon)} \times \left(\tilde{\alpha} - \frac{\alpha_2 m_1}{\rho} \right) \times \rho^3 \mathcal{I}_2(u) - \frac{\beta\kappa_\infty \alpha_2 \nu^2 m_1^2}{4\rho^4(1 - \epsilon)} \times \left[\rho^3 \mathcal{I}_2(u) + \frac{1}{3} \rho^4 \mathcal{I}_3(u) + \frac{1}{24} \rho^5 \mathcal{I}_4(u) \right], \end{aligned} \quad (38)$$

where, for $p \in \mathbb{N}$ and $u \geq 0$, $\mathcal{I}_p(u) := \exp(2\iota_c u) \int_0^u s^p \exp(-2\iota_c s) ds$, and $\tilde{\alpha}, \alpha_2$ are defined in (26).

B.4 Determination of the optimal strategy

The final step of the proof is to determine the strategy X^* such that $\Pi(X^*)$ is a martingale, or equivalently such that A^{X^*} is constant. Equations (25) and (27) yield

$$\begin{aligned} dA_t^{X^*} &= \frac{\rho}{1-\epsilon} dt \times \left[\frac{D_t - (1-\epsilon)X_t}{2 + \rho(T-t)} - D_t + k(T-t) \delta_t \right]^2 \\ &= \frac{\rho/(1-\epsilon)}{[2 + \rho(T-t)]^2} dt \times \left[(1-\epsilon)X_t + [1 + \rho(T-t)] D_t - [2 + \rho(T-t)] k(T-t) \delta_t \right]^2. \end{aligned}$$

Thus, A^{X^*} is constant on $(0, T)$ if, and only if

$$\text{a.s. , dt-a.e. on } (0, T) \text{ , } (1-\epsilon)X_t^* = -[1 + \rho(T-t)] D_t^* + [2 + \rho(T-t)] k(T-t) \delta_t, \quad (40)$$

where $D = D^*$ when the strategy X^* is used by the strategic trader. Then, we characterize the strategy X^* on $[0, T]$ with the three following steps:

- The initial jump ΔX_0^* of the strategy is such that (X^*, D^*) satisfies equation (40) at time $t = 0^+$.
- The strategy X^* on $(0, T)$ is obtained by differentiating equation (40).
- The final jump $\Delta X_T^* = -X_T^*$ closes the position of the strategic trader at time T .

We need the following lemma in the sequel.

Lemma B.1. *Let $\phi : [0, T] \rightarrow \mathbb{R}$ be a measurable function, and for $0 \leq s \leq t \leq T$, $\Phi(s, t) := \int_s^t \phi(u) \exp(-\beta u) du$. We then have for all $t \in [0, T]$*

$$\int_0^t \phi(u) \delta_u du = \delta_0 \Phi(0, t) + \Theta_{\chi_t} \Phi(\tau_{\chi_t}, t) + \sum_{i=1}^{\chi_t-1} \Theta_i \Phi(\tau_i, \tau_{i+1})$$

Proof. The proof is straightforward since for $u \in [\chi_t, t]$, $\delta_u = \delta_0 \exp(-\beta u) + \exp(-\beta u) \Theta_{\chi_t}$ and for $i \in \{0, \dots, \chi_t - 1\}$ and $u \in [\tau_i, \tau_{i+1})$, $\delta_u = \delta_0 \exp(-\beta u) + \exp(-\beta u) \Theta_i$. \square

To determine the optimal strategy, only the function k given by (36) comes into play, thus the cases $\eta = 0$ and $\eta \neq 0$ can be treated simultaneously. We also note that

$$\frac{d}{du}[u^2 \omega(\eta u)] = u\zeta(\eta u) \quad \text{and} \quad \frac{d}{du}[u \zeta(\eta u)] = \exp(-\eta u)$$

hold for for all $u \geq 0$ and $\eta \in \mathbb{R}$. We use Equations (36) and (40) to obtain the following characterization of the strategy X^* : a.s., dt-a.e. on $(0, T)$,

$$(1-\epsilon)X_t^* = -[1 + \rho(T-t)] D_t^* + \frac{m_1}{2\rho} \times [2 + \rho(T-t) \times \{1 + \zeta(\eta(T-t)) + \nu\rho(T-t) \omega(\eta(T-t))\}] \delta_t. \quad (41)$$

The initial jump of X^* at $t = 0$ is such that (41) is verified for $t = 0^+$:

$$(1-\epsilon)(x_0 + \Delta X_0^*) = -[1 + \rho T] (D_0 + (1-\epsilon)\Delta X_0^*) + \frac{m_1}{2\rho} \times [2 + \rho T \times \{1 + \zeta(\eta T) + \nu\rho T \omega(\eta T)\}] \delta_0, \quad (42)$$

which gives the initial trade at time 0 as given in Appendix A.

We differentiate Equation (41) to get

$$(1 - \epsilon)dX_t^* = \rho D_t^* dt - [1 + \rho(T - t)] dD_t^* - \frac{m_1}{2} \times [1 + \exp(-\eta(T - t)) + \nu\rho(T - t)\zeta(\eta(T - t))] \delta_t dt + \frac{m_1}{2\rho} \times [2 + \rho(T - t) \times \{1 + \zeta(\eta(T - t)) + \nu\rho(T - t) \omega(\eta(T - t))\}] d\delta_t.$$

This yields, using $d\delta_t = -\beta \delta_t dt + dI_t$,

$$(1 - \epsilon)dX_t^* = \rho D_t^* dt - m_1 \phi_\eta(t) \delta_t dt + \frac{1 + \rho(T - t)}{2 + \rho(T - t)} \left\{ \frac{m_1}{\rho} dI_t - (1 - \nu) dN_t \right\} + \frac{m_1}{2\rho} \times \frac{\rho(T - t) \times \{\zeta(\eta(T - t)) - 1 + \nu\rho(T - t) \omega(\eta(T - t))\}}{2 + \rho(T - t)} dI_t, \quad (43)$$

where for $t \in [0, T]$

$$\phi_\eta(t) := \frac{1}{2} \times \frac{1 + \exp(-\eta(T - t)) + \nu\rho(T - t)\zeta(\eta(T - t)) + \frac{\beta}{\rho} [2 + \rho(T - t) \times \{1 + \zeta(\eta(T - t)) + \nu\rho(T - t) \omega(\eta(T - t))\}]}{2 + \rho(T - t)}$$

and $\delta_t = \delta_0 \exp(-\beta t) + \sum_{0 < \tau \leq t} \exp(-\beta(t - \tau)) \Delta I_\tau$. For $t \in (0, T)$,

$$\begin{aligned} dD_t^* &= -\rho D_t^* dt + (1 - \epsilon)dX_t^* + (1 - \nu)dN_t \\ &= -m_1 \phi_\eta(t) \delta_t dt + \frac{(1 - \nu) dN_t}{2 + \rho(T - t)} + \frac{m_1}{2\rho} \times \frac{2 + \rho(T - t) \times \{1 + \zeta(\eta(T - t)) + \nu\rho(T - t) \omega(\eta(T - t))\}}{2 + \rho(T - t)} dI_t, \end{aligned}$$

and we have

$$\begin{aligned} D_{0+}^* &= D_0 + (1 - \epsilon)\Delta X_0^* = \frac{D_0 - (1 - \epsilon)x_0}{2 + \rho T} + \frac{m_1}{2\rho} \times \frac{2 + \rho T \times \{1 + \zeta(\eta T) + \nu\rho T \omega(\eta T)\}}{2 + \rho T} \delta_0 \\ \int_{(0,t]} dD_u^* &= -m_1 \int_{(0,t]} \phi_\eta(u) \delta_u du + \sum_{0 < \tau \leq t} \frac{(1 - \nu) \Delta N_\tau}{2 + \rho(T - \tau)} \\ &\quad + \frac{m_1}{2\rho} \times \sum_{0 < \tau \leq t} \frac{2 + \rho(T - \tau) \times \{1 + \zeta(\eta(T - \tau)) + \nu\rho(T - \tau) \omega(\eta(T - \tau))\}}{2 + \rho(T - \tau)} \Delta I_\tau. \end{aligned}$$

We define $\Phi_\eta(s, t) := \int_s^t \phi_\eta(u) \exp(-\beta u) du$ for $0 \leq s \leq t \leq T$. Lemma B.1 yields for $t \in [0, T]$

$$\int_0^t \phi_\eta(u) \delta_u du = \delta_0 \Phi_\eta(0, t) + \Theta_{\chi_t} \Phi_\eta(\tau_{\chi_t}, t) + \sum_{i=1}^{\chi_t-1} \Theta_i \Phi_\eta(\tau_i, \tau_{i+1}).$$

We obtain the expression of D_t^* for $t \in (0, T)$

$$\begin{aligned} D_t^* &= \frac{D_0 - (1 - \epsilon)x_0}{2 + \rho T} + \frac{\delta_0 m_1}{2\rho} \times \left[\frac{2 + \rho T \times \{1 + \zeta(\eta T) + \nu\rho T \omega(\eta T)\}}{2 + \rho T} - 2\rho \Phi_\eta(0, t) \right] \\ &\quad - m_1 \left[\Theta_{\chi_t} \Phi_\eta(\tau_{\chi_t}, t) + \sum_{i=1}^{\chi_t-1} \Theta_i \Phi_\eta(\tau_i, \tau_{i+1}) \right] + \sum_{0 < \tau \leq t} \frac{(1 - \nu) \Delta N_\tau}{2 + \rho(T - \tau)} \\ &\quad + \frac{m_1}{2\rho} \times \sum_{0 < \tau \leq t} \frac{2 + \rho(T - \tau) \times \{1 + \zeta(\eta(T - \tau)) + \nu\rho(T - \tau) \omega(\eta(T - \tau))\}}{2 + \rho(T - \tau)} \Delta I_\tau. \end{aligned}$$

From (43), the strategy X^* on $(0, T)$ is as given in Appendix A. By using again (41), we also get the final trade at time T .

We determine the function Φ_η in the case $\eta \neq 0$ (similar and simpler calculations yield the result for $\eta = 0$). We write

$$\begin{aligned}\exp(-\eta(T-t)) \times \exp(-\beta t) &= \exp(-\beta T) \times \exp(\alpha(T-t)), \\ (T-t)\zeta(\eta(T-t)) \times \exp(-\beta t) &= \frac{\exp(-\beta T)}{\eta} \times [\exp(\beta(T-t)) - \exp(\alpha(T-t))].\end{aligned}$$

Thus, $\phi_\eta(t) \times \exp(\beta(T-t))$ is equal to

$$\frac{\beta}{2} \left(\frac{1}{\rho} + \frac{\nu}{\eta} \right) \times \exp(\beta(T-t)) + \left[\frac{1}{2} + \frac{\nu(\rho-2\beta)}{2\eta} + \frac{\beta}{2\eta} \left(1 - \frac{\nu\rho}{\eta} \right) \right] \frac{\exp(\beta(T-t))}{2 + \rho(T-t)} + \left[\frac{1}{2} - \frac{\nu\rho}{2\eta} - \frac{\beta}{2\eta} \left(1 - \frac{\nu\rho}{\eta} \right) \right] \frac{\exp(\alpha(T-t))}{2 + \rho(T-t)},$$

which yields for $0 \leq s \leq t \leq T$,

$$\begin{aligned}\Phi_\eta(s, t) &= \frac{1}{2} \left(\frac{1}{\rho} + \frac{\nu}{\eta} \right) \times [\exp(-\beta s) - \exp(-\beta t)] \\ &\quad + \frac{\exp(-\beta T)}{2\rho} \times \left[1 + \frac{\nu(\rho-2\beta)}{\eta} + \frac{\beta}{\eta} \left(1 - \frac{\nu\rho}{\eta} \right) \right] \times [L(\rho, \beta, T-s) - L(\rho, \beta, T-t)] \\ &\quad + \frac{\exp(-\beta T)}{2\rho} \times \left[1 - \frac{\nu\rho}{\eta} - \frac{\beta}{\eta} \left(1 - \frac{\nu\rho}{\eta} \right) \right] \times [L(\rho, \alpha, T-s) - L(\rho, \alpha, T-t)].\end{aligned}$$

with $\eta = \beta - \alpha \neq 0$.

C Proof of Theorem 2.1

Let X be an admissible strategy. We introduce the following processes: $S_t^N = S_0 + \frac{\nu}{q}(N_t - N_0)$, $S_t^X = \frac{\epsilon}{q}(X_t - X_0)$,

$$dD_t^N = -\rho D_t^N dt + \frac{1-\nu}{q} dN_t \text{ and } dD_t^X = -\rho D_t^X dt + \frac{1-\epsilon}{q} dX_t,$$

with $D_0^N = D_0$ and $D_0^X = 0$. Thus, we have $S = S^N + S^X$, $D = D^N + D^X$ and thus $P = P^N + P^X$, where $P^N = S^N + D^N$ and $P^X = S^X + D^X$. From (4), we have

$$C(X) = \int_{[0, T)} P_u^N dX_u - P_T^N X_T + C^{\text{OW}}(X),$$

where

$$C^{\text{OW}}(X) = \int_{[0, T)} P_u^X dX_u + \frac{1}{2q} \sum_{0 \leq \tau < T} (\Delta X_\tau)^2 - P_T^X X_T + \frac{1}{2q} X_T^2$$

is a deterministic function of X that corresponds to the cost when $N \equiv 0$, which is the Obizhaeva and Wang model. We now make an integration by parts as in Remark 2.6 and get that

$$\int_{[0, T)} P_u^N dX_u - P_T^N X_T = - \int_{[0, T)} X_u dP_u^N.$$

When P^N is a martingale, this term has a null expectation. Therefore, the optimal execution strategy is the same as in the Obizhaeva and Wang model, see Gatheral, Schied and Slynko [23], Example 2.12, and there is no PMS. Otherwise, we can find $0 \leq s < t \leq T$ such that $\mathbb{E}[P_t^N | \mathcal{F}_s]$ and P_s^N are not almost surely equal. In this case, we consider the strategy $X_u = \mathbb{E}[P_t^N - P_s^N | \mathcal{F}_s] \mathbf{1}_{u \in (s, t]}$ that is a round-trip, i.e. $X_0 = X_{T+} = 0$. We then get

$$\mathbb{E} \left[- \int_{[0, T)} X_u dP_u^N \right] = -\mathbb{E}[(P_t^N - P_s^N) \mathbb{E}[P_t^N - P_s^N | \mathcal{F}_s]] = -\mathbb{E}[\mathbb{E}[P_t^N - P_s^N | \mathcal{F}_s]^2] < 0.$$

Since $C^{\text{OW}}(cX) = c^2 C^{\text{OW}}(X)$, we can find c small enough such that $E[C(cX)] = -c \mathbb{E}[\mathbb{E}[P_t^N - P_s^N | \mathcal{F}_s]^2] + c^2 C^{\text{OW}}(X) < 0$, and therefore cX is a PMS.

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